

Implementation of an iterative ensemble smoother for big-data assimilation and reservoir history matching

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Parameter-estimation problem

Nonlinear model and measurements

$$\mathbf{y} = \mathbf{g}(\mathbf{x}) \quad \mathbf{d} \leftarrow \mathbf{y} + \mathbf{e}$$

Bayesian formulation

$$f(\mathbf{x}, \mathbf{y} | \mathbf{d}) \propto f(\mathbf{d} | \mathbf{y}) f(\mathbf{y} | \mathbf{x}) f(\mathbf{x})$$

- ▶ Standard History-Matching problem for oil-reservoir models.

“Indirect” DA update

Nonlinear model

$$\mathbf{x}_{i+1} = \mathbf{m}(\mathbf{x}_i)$$

Nonlinear measurement operator and measurements

$$\mathbf{y} = \mathbf{h}(\mathbf{x}_{i+1}) = \mathbf{h}(\mathbf{m}(\mathbf{x}_i)) = \mathbf{g}(\mathbf{x}_i), \quad \mathbf{d} \leftarrow \mathbf{y} + \mathbf{e}$$

Bayesian formulation

$$f(\mathbf{x}_i, \mathbf{y} | \mathbf{d}) \propto f(\mathbf{d} | \mathbf{y}) f(\mathbf{y} | \mathbf{x}_i) f(\mathbf{x}_i)$$

- ▶ Smoother update step in sequential data assimilation

Marginal pdf

Nonlinear model and measurements

$$\mathbf{y} = \mathbf{g}(\mathbf{x}) \quad \mathbf{d} \leftarrow \mathbf{y} + \mathbf{e}$$

Model pdf

$$f(\mathbf{y}|\mathbf{x}) = \delta(\mathbf{y} - \mathbf{g}(\mathbf{x}))$$

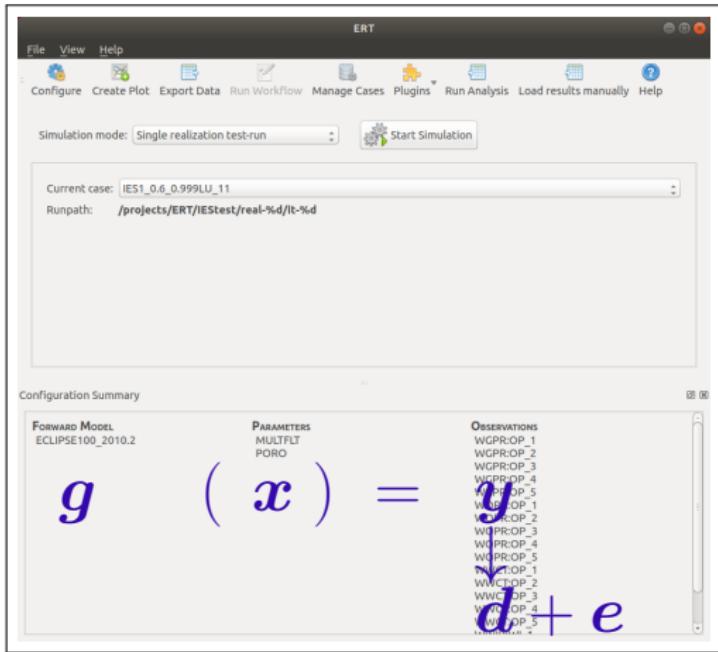
Bayesian formulation

$$f(\mathbf{x}, \mathbf{y}|\mathbf{d}) \propto f(\mathbf{d}|\mathbf{y})f(\mathbf{y}|\mathbf{x})f(\mathbf{x})$$

Marginal pdf

$$f(\mathbf{x}|\mathbf{d}) \propto \int f(\mathbf{d}|\mathbf{y})f(\mathbf{y}|\mathbf{x})f(\mathbf{x})d\mathbf{y} = f(\mathbf{d}|\mathbf{g}(\mathbf{x}))f(\mathbf{x})$$

ERT



Gaussian priors

Maximizing $f(\mathbf{x}|\mathbf{d})$ is equivalent to minimizing

$$\mathcal{J}(\mathbf{x}) = (\mathbf{x} - \mathbf{x}^f)^T \mathbf{C}_{xx}^{-1} (\mathbf{x} - \mathbf{x}^f) + (\mathbf{g}(\mathbf{x}) - \mathbf{d})^T \mathbf{C}_{dd}^{-1} (\mathbf{g}(\mathbf{x}) - \mathbf{d}).$$

Exact direct solution in the case when $\mathbf{g}(\mathbf{x})$ is linear (KF update)

$$\mathbf{x}^a = \mathbf{x}^f + \mathbf{K}(\mathbf{d} - \mathbf{g}(\mathbf{x})), \quad \mathbf{C}_{xx}^a = (\mathbf{I} - \mathbf{K}\mathbf{H})\mathbf{C}_{xx}^f.$$

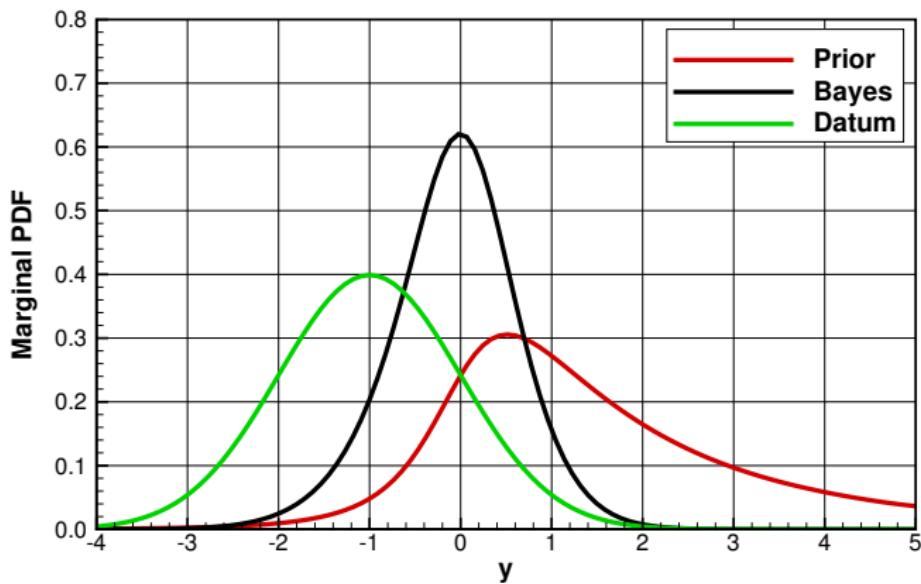
Ensemble representation (EnKF)

$$\mathbf{x}_j^a = \mathbf{x}_j^f + \overline{\mathbf{K}}(\mathbf{d}_j - \mathbf{g}(\mathbf{x}_j)), \quad \overline{\mathbf{x}}^a \rightarrow \mathbf{x}^a, \quad \overline{\mathbf{C}}_{xx}^a \rightarrow \mathbf{C}_{xx}^a.$$

Equivalent to minimizing (linear case)

$$\mathcal{J}(\mathbf{x}_j) = (\mathbf{x}_j - \mathbf{x}_j^f)^T \mathbf{C}_{xx}^{-1} (\mathbf{x}_j - \mathbf{x}_j^f) + (\mathbf{g}(\mathbf{x}_j) - \mathbf{d}_j)^T \mathbf{C}_{dd}^{-1} (\mathbf{g}(\mathbf{x}_j) - \mathbf{d}_j).$$

Bayesian update



EnKF update

Prior ensemble

$$\mathbf{y}_j^{\text{f}} = \mathbf{g}(\mathbf{x}_j^{\text{f}})$$

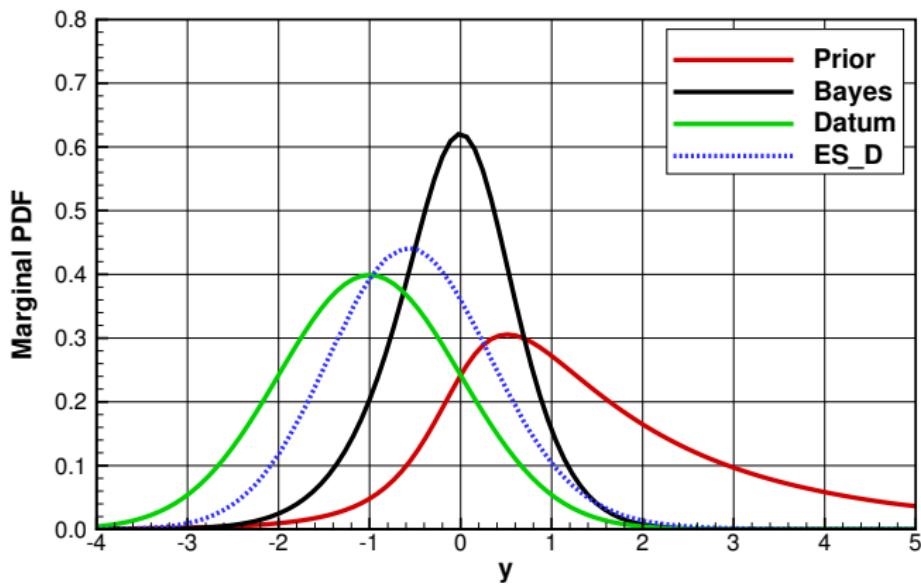
Ensemble of measurements of \mathbf{y}

$$\mathbf{d}_j = \mathbf{d} + \mathbf{e}_j$$

Direct update of \mathbf{y}_j

$$\mathbf{y}_j^{\text{a}} = \mathbf{y}_j^{\text{f}} + \overline{\mathbf{C}}_{yy}^{\text{f}} (\overline{\mathbf{C}}_{yy}^{\text{f}} + \mathbf{C}_{dd})^{-1} (\mathbf{d}_j - \mathbf{y}_j^{\text{f}})$$

EnKF (direct) update



Ensemble smoother (ES) indirect update

Prior

$$\mathbf{y}_j^{\text{f}} = \mathbf{g}(\mathbf{x}_j^{\text{f}})$$

Ensemble of measurements of \mathbf{y}

$$\mathbf{d}_j = \mathbf{d} + \mathbf{e}_j$$

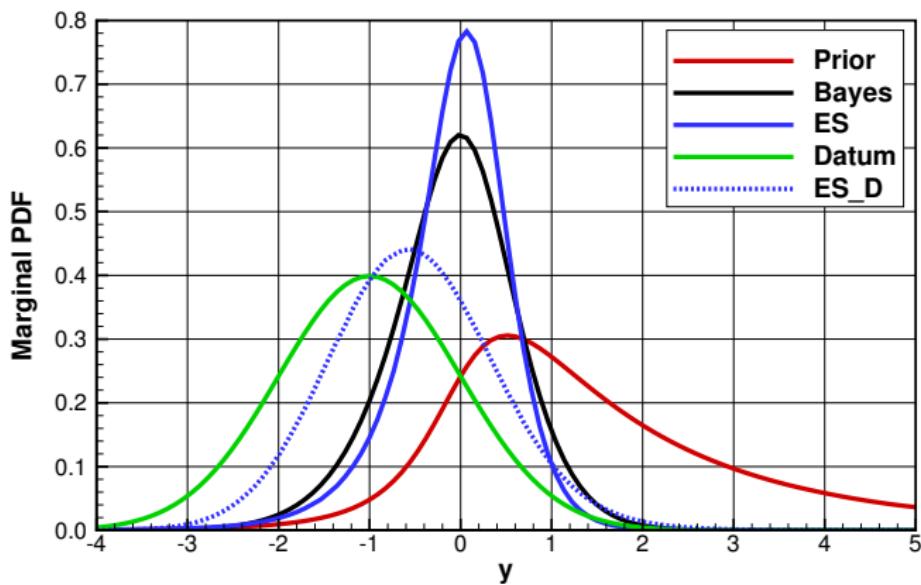
Smoother update of \mathbf{x}

$$\mathbf{x}_j^{\text{a}} = \mathbf{x}_j^{\text{f}} + \overline{\mathbf{C}}_{xy}^{\text{f}} (\overline{\mathbf{C}}_{yy}^{\text{f}} + \mathbf{C}_{dd})^{-1} (\mathbf{d}_j - \mathbf{y}_j^{\text{f}})$$

Indirect update of \mathbf{y}

$$\mathbf{y}_j^{\text{a}} = \mathbf{g}(\mathbf{x}_j^{\text{a}})$$

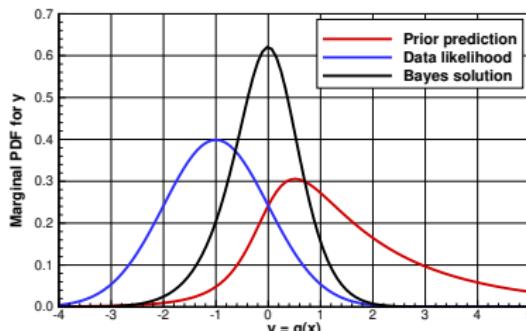
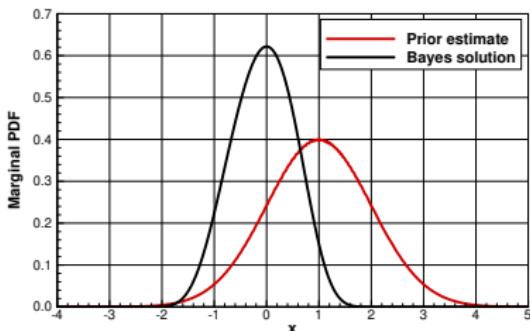
ES (indirect) update



History matching problem

Bayes theorem gives posterior probability function for parameters x

$$f(x|d) \propto f(d|g(x))f(x)$$



- ▶ x represents static parameters like porosity, permeability, ...
- ▶ y could represent predicted OPR from Eclipse.
- ▶ Prior pdf represents uncertainty of x .
- ▶ Prior prediction pdf represents uncertainty of $y = g(x)$.
- ▶ Data likelihood represents uncertainty of measurement d .

ES and IES

Gaussian assumption!

Approximate sampling of posterior pdf

$$f(\mathbf{x}|\mathbf{d}) \propto f(\mathbf{d}|\mathbf{g}(\mathbf{x}))f(\mathbf{x})$$

by minimizing an ensemble of cost functions

$$\mathcal{J}(\mathbf{x}_j) = (\mathbf{x}_j - \mathbf{x}_j^{\text{f}})^T \mathbf{C}_{xx}^{-1} (\mathbf{x}_j - \mathbf{x}_j^{\text{f}}) + (\mathbf{g}(\mathbf{x}_j) - \mathbf{d}_j)^T \mathbf{C}_{dd}^{-1} (\mathbf{g}(\mathbf{x}_j) - \mathbf{d}_j).$$

Prior misfit data misfit

Solved by

1. Ensemble Smoother
2. Iterative Ensemble Smoother

Ensemble Smoother

Approximately solves $\nabla \mathcal{J} = 0$

$$\mathbf{C}_{xx}^{-1}(\mathbf{x}_j - \mathbf{x}_j^f) + \nabla_x \mathbf{g}(\mathbf{x}_j) \mathbf{C}_{dd}^{-1}(\mathbf{g}(\mathbf{x}_j) - \mathbf{d}_j) = 0.$$

- ▶ Define $\mathbf{G}_j^T = \nabla_x \mathbf{g}(\mathbf{x}_j)$.
- ▶ ES applies a linearization $\mathbf{g}(\mathbf{x}_j) = \mathbf{x}_j^f + \mathbf{G}_j(\mathbf{x}_j - \mathbf{x}_j^f)$.
- ▶ Model sensitivities replaced by least-squares fit $\mathbf{C}_{yx} = \mathbf{G} \mathbf{C}_{xx}$
- ▶ ES uses ensemble covariances $\overline{\mathbf{C}}_{xy}$, $\overline{\mathbf{C}}_{xx}$, and $\overline{\mathbf{C}}_{dd}$

$$\mathbf{x}_j^f \leftarrow \mathcal{N}(\mathbf{x}^f, \mathbf{C}_{xx}^f), \quad \mathbf{d}_j \leftarrow \mathcal{N}(\mathbf{d}, \mathbf{C}_{dd}),$$

$$\mathbf{y}_j^f = \mathbf{g}(\mathbf{x}_j^f),$$

$$\mathbf{x}_j^a = \mathbf{x}_j^f + \overline{\mathbf{C}}_{xy} \left(\overline{\mathbf{C}}_{yx} (\overline{\mathbf{C}}_{xx})^{-1} \overline{\mathbf{C}}_{xy} + \overline{\mathbf{C}}_{dd} \right)^{-1} \left(\mathbf{d}_j - \mathbf{y}_j^f \right),$$

$$\mathbf{y}_j^a = \mathbf{g}(\mathbf{x}_j^a).$$

Iterative Ensemble Smoother

State-space formulation with Gauss-Newton iterations

$$\mathcal{J}(\mathbf{x}_j) = (\mathbf{x}_j - \mathbf{x}_j^f)^T \mathbf{C}_{xx}^{-1} (\mathbf{x}_j - \mathbf{x}_j^f) + (\mathbf{g}(\mathbf{x}_j) - \mathbf{d}_j)^T \mathbf{C}_{dd}^{-1} (\mathbf{g}(\mathbf{x}_j) - \mathbf{d}_j)$$

with gradient and Hessian

$$\nabla_{\mathbf{x}} \mathcal{J} = \mathbf{C}_{xx}^{-1} (\mathbf{x}_j - \mathbf{x}_j^f) + \nabla_{\mathbf{x}} \mathbf{g}(\mathbf{x}_j) \mathbf{C}_{dd}^{-1} (\mathbf{g}(\mathbf{x}_j) - \mathbf{d}_j),$$

$$\nabla_{\mathbf{x}} \nabla_{\mathbf{x}} \mathcal{J} \approx \mathbf{C}_{xx}^{-1} + \nabla_{\mathbf{x}} \mathbf{g}(\mathbf{x}_j) \mathbf{C}_{dd}^{-1} (\nabla_{\mathbf{x}} \mathbf{g}(\mathbf{x}_j))^T$$

Iterate

$$\mathbf{x}_j^{i+1} = \mathbf{x}_j^i - \gamma (\nabla \nabla \mathcal{J}_i)^{-1} \nabla \mathcal{J}_j^i$$

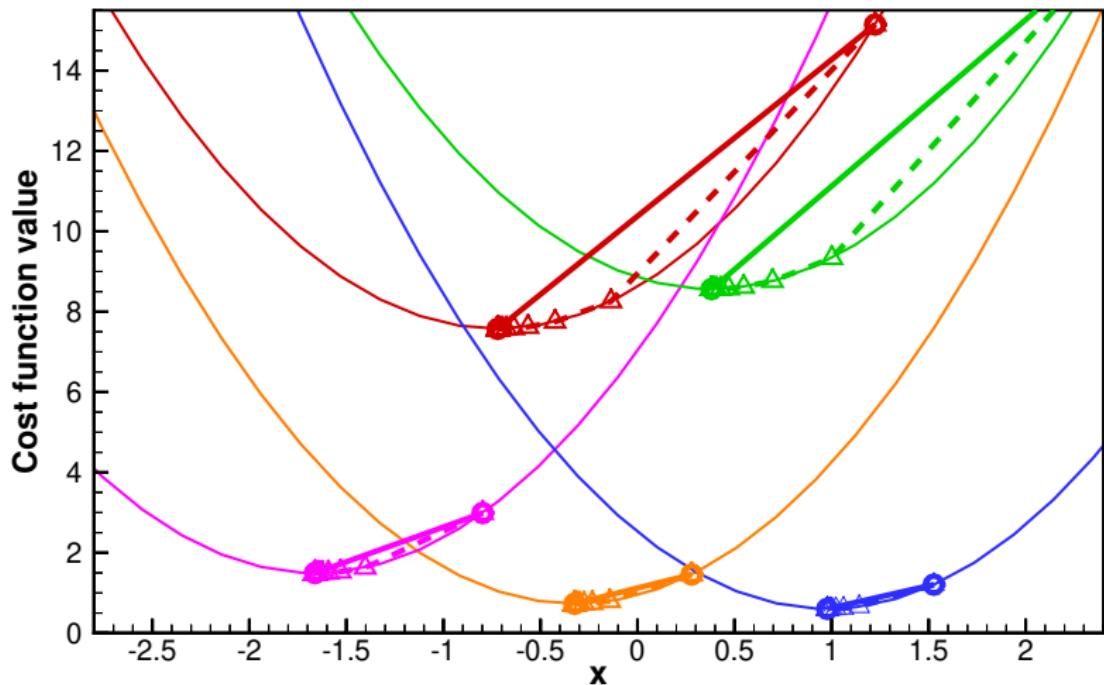
$$\mathbf{y}_j^{i+1} = \mathbf{g}(\mathbf{x}_j^{i+1})$$

IES gradient

$$\begin{aligned} (\nabla \nabla \mathcal{J}_i)^{-1} \nabla \mathcal{J}_{j,i} &= \bar{\mathbf{C}}_{xx}^i \bar{\mathbf{C}}_{xx}^{-1} (\mathbf{x}_{j,i} - \mathbf{x}_j^f) \\ &\quad - \bar{\mathbf{C}}_{xy}^i \left(\bar{\mathbf{C}}_{yx}^i (\bar{\mathbf{C}}_{xx}^i)^{-1} \bar{\mathbf{C}}_{xy}^i + \bar{\mathbf{C}}_{dd} \right)^{-1} \\ &\quad \times \left(\bar{\mathbf{C}}_{yx}^i \bar{\mathbf{C}}_{xx}^{-1} (\mathbf{x}_{j,i} - \mathbf{x}_j^f) - (\mathbf{g}(\mathbf{x}_{j,i}) - \mathbf{d}_j) \right). \end{aligned}$$

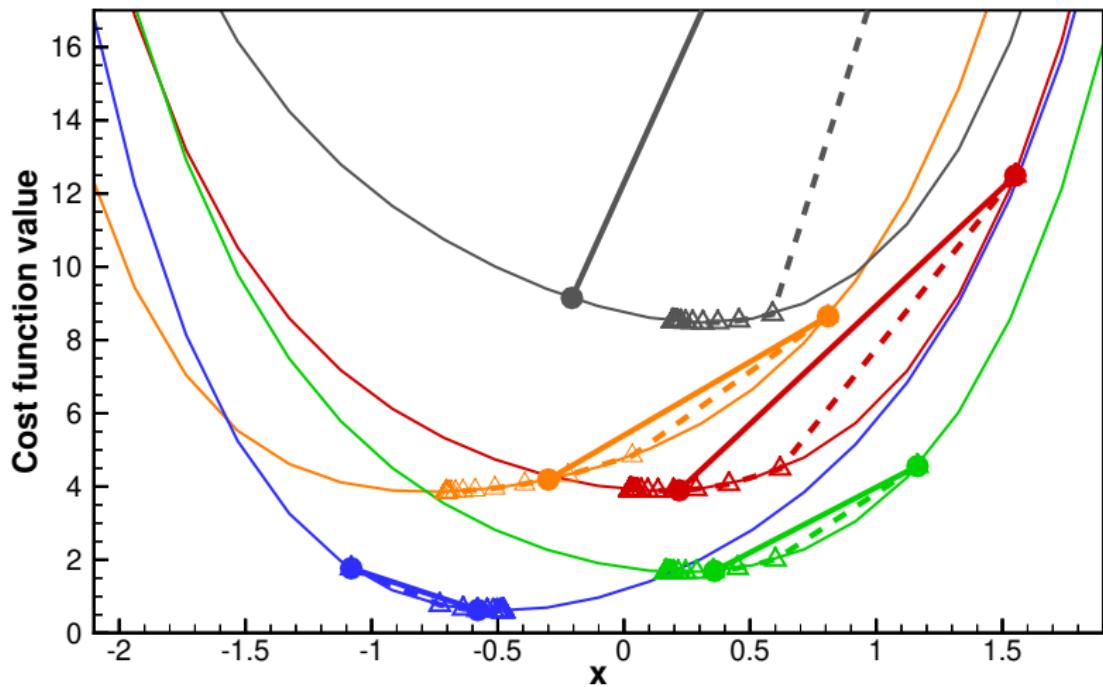
- ▶ Model sensitivities replaced with averaged least squares fit.
- ▶ Covariances replaced with ensemble representations.
- ▶ Leftmost factor of gradient is ensemble anomaly matrix.
- ▶ Solution is searched for in the ensemble subspace

ES and IES illustration: Linear model



- ▶ IES and ES both finds global minimum.
- ▶ Samples exactly posterior pdf.

ES and IES illustration: Non-linear model



- ▶ IES gets closer to minimum than ES
- ▶ Approximate sampling of posterior pdf.

ESMDA: Rewriting likelihood

Approximate sampling of

$$f(\mathbf{x}|\mathbf{d}) \propto f(\mathbf{d}|\mathbf{g}(\mathbf{x})) f(\mathbf{x})$$

by gradually introducing the measurements (?)

$$\begin{aligned} f(\mathbf{x}|\mathbf{d}) &= f(\mathbf{d}|\mathbf{y})^{\frac{1}{\alpha_N}} \cdots f(\mathbf{d}|\mathbf{y})^{\frac{1}{\alpha_2}} f(\mathbf{d}|\mathbf{y})^{\frac{1}{\alpha_1}} f(\mathbf{x}) \\ &= f(\mathbf{d}|\mathbf{y})^{\left(\sum_{i=1}^N \frac{1}{\alpha_i}\right)} f(\mathbf{x}) \\ &= f(\mathbf{d}|\mathbf{y}) f(\mathbf{x}) \end{aligned}$$

with

$$\sum_{i=1}^N \frac{1}{\alpha_i} = 1$$

Some definitions

Prior ensemble and perturbed measurements

$$\mathbf{X} = \left(\mathbf{x}_1^f, \mathbf{x}_2^f, \dots, \mathbf{x}_N^f \right) \quad \text{and} \quad \mathbf{D} = \left(\mathbf{d}_1, \mathbf{d}_2, \dots, \mathbf{d}_N \right)$$

Ensemble means

$$\bar{\mathbf{x}} = \frac{1}{N} \sum_{j=1}^N \mathbf{x}_j \quad \text{and} \quad \bar{\mathbf{d}} = \frac{1}{N} \sum_{j=1}^N \mathbf{d}_j$$

Ensemble anomaly matrices and covariances

$$\mathbf{A} = \mathbf{X} \left(\mathbf{I}_N - \frac{1}{N} \mathbf{1} \mathbf{1}^T \right) / \sqrt{N-1} \quad \rightarrow \quad \bar{\mathbf{C}}_{xx} = \mathbf{A} \mathbf{A}^T$$

$$\mathbf{E} = \mathbf{D} \left(\mathbf{I}_N - \frac{1}{N} \mathbf{1} \mathbf{1}^T \right) / \sqrt{N-1} \quad \rightarrow \quad \bar{\mathbf{C}}_{dd} = \mathbf{E} \mathbf{E}^T$$

IES: Ensemble subspace version

Original cost functions

$$\mathcal{J}(\mathbf{x}_j) = (\mathbf{x}_j - \mathbf{x}_j^f)^T \mathbf{C}_{xx}^{-1} (\mathbf{x}_j - \mathbf{x}_j^f) + (\mathbf{g}(\mathbf{x}_j) - \mathbf{d}_j)^T \mathbf{C}_{dd}^{-1} (\mathbf{g}(\mathbf{x}_j) - \mathbf{d}_j).$$

Solution is contained in the ensemble subspace, thus

$$\mathbf{x}_j^a = \mathbf{x}_j^f + \mathbf{A}\mathbf{w}_j,$$

and,

$$\mathcal{J}(\mathbf{w}_j) = \mathbf{w}_j^T \mathbf{w}_j + \left(\mathbf{g}(\mathbf{x}_j^f + \mathbf{A}\mathbf{w}_j) - \mathbf{d}_j \right)^T \mathbf{C}_{dd}^{-1} \left(\mathbf{g}(\mathbf{x}_j^f + \mathbf{A}\mathbf{w}_j) - \mathbf{d}_j \right)$$

Reduces dimension of problem from state size to ensemble size.

$$\mathbf{w}_j^{i+1} = \mathbf{w}_j^i - \gamma \nabla \mathcal{J}_j^i$$

Gradient and Hessian of cost function

Gradient

$$\nabla \mathcal{J}(\boldsymbol{w}_j) = \mathbf{2}\boldsymbol{w}_j + 2(\mathbf{G}_j\mathbf{A})^T \mathbf{C}_{dd}^{-1} (\mathbf{g}(\mathbf{x}_j^f + \mathbf{A}\boldsymbol{w}_j) - \mathbf{d}_j),$$

Hessian (approximate)

$$\nabla\nabla \mathcal{J}(\boldsymbol{w}_j) \approx \mathbf{2}\mathbf{I} + 2(\mathbf{G}_j\mathbf{A})^T \mathbf{C}_{dd}^{-1} (\mathbf{G}_j\mathbf{A})$$

Gauss-Newton iterations

$$\mathbf{w}_j^{i+1} = \mathbf{w}_j^i - \gamma \Delta \mathbf{w}_j^i$$

$$\begin{aligned}\Delta \mathbf{w}_j^i &= \left\{ \mathbf{w}_j^i - (\mathbf{G}_j^i \mathbf{A})^T \left((\mathbf{G}_j^i \mathbf{A})(\mathbf{G}_j^i \mathbf{A})^T + \mathbf{C}_{dd} \right)^{-1} \right. \\ &\quad \times \left. \left((\mathbf{G}_j^i \mathbf{A}) \mathbf{w}_j^i + \mathbf{d}_j - \mathbf{g}(\mathbf{x}_j^f + \mathbf{A} \mathbf{w}_j^i) \right) \right\}.\end{aligned}$$

with

$$\mathbf{G}_j^i = (\nabla \mathbf{g} |_{\mathbf{x}_j^f + \mathbf{A} \mathbf{w}_j^i})^T.$$

$G_j^i A$

Replace G_j^i with average sensitivity G^i

$G_j^i A$

Replace G_j^i with average sensitivity G^i

Define the linear regression

$$\bar{G}_i = Y_i A_i^+$$

$$Y_i = g(X_i) \left(I - \frac{1}{N} \mathbf{1} \mathbf{1}^T \right) \Big/ \sqrt{N-1}$$

$$\mathbf{G}_j^i \mathbf{A}$$

Replace \mathbf{G}_j^i with average sensitivity $\bar{\mathbf{G}}^i$

Define the linear regression

$$\bar{\mathbf{G}}_i = \mathbf{Y}_i \mathbf{A}_i^+$$

Write

$$\mathbf{G}_j^i \mathbf{A} \approx \bar{\mathbf{G}}^i \mathbf{A} = \mathbf{Y}_i \mathbf{A}_i^+ \mathbf{A}$$

$$\mathbf{Y}_i = \mathbf{g}(\mathbf{X}_i) \left(\mathbf{I} - \frac{1}{N} \mathbf{1} \mathbf{1}^T \right) \Big/ \sqrt{N-1}$$

$$\mathbf{G}_j^i \mathbf{A}$$

Replace \mathbf{G}_j^i with average sensitivity $\bar{\mathbf{G}}^i$

Define the linear regression

$$\bar{\mathbf{G}}_i = \mathbf{Y}_i \mathbf{A}_i^+$$

Write

$$\begin{aligned}\mathbf{G}_j^i \mathbf{A} &\approx \bar{\mathbf{G}}^i \mathbf{A} = \mathbf{Y}_i \mathbf{A}_i^+ \mathbf{A} \\ &= \mathbf{Y}_i \mathbf{A}_i^+ \mathbf{A}_i \boldsymbol{\Omega}_i^{-1} \quad \mathbf{A} = \mathbf{A}_i \boldsymbol{\Omega}_i^{-1}\end{aligned}$$

$$\mathbf{Y}_i = \mathbf{g}(\mathbf{X}_i) \left(\mathbf{I} - \frac{1}{N} \mathbf{1} \mathbf{1}^T \right) \Big/ \sqrt{N-1}$$

$$\boldsymbol{\Omega}_i = \mathbf{I} + \mathbf{W}_i \left(\mathbf{I} - \frac{1}{N} \mathbf{1} \mathbf{1}^T \right) \Big/ \sqrt{N-1}$$

$$\mathbf{G}_j^i \mathbf{A}$$

Replace \mathbf{G}_j^i with average sensitivity $\bar{\mathbf{G}}^i$

Define the linear regression

$$\bar{\mathbf{G}}_i = \mathbf{Y}_i \mathbf{A}_i^+$$

Write

$$\begin{aligned}\mathbf{G}_j^i \mathbf{A} &\approx \bar{\mathbf{G}}^i \mathbf{A} = \mathbf{Y}_i \mathbf{A}_i^+ \mathbf{A} \\ &= \mathbf{Y}_i \mathbf{A}_i^+ \mathbf{A}_i \boldsymbol{\Omega}_i^{-1} & \mathbf{A} = \mathbf{A}_i \boldsymbol{\Omega}_i^{-1} \\ &= \mathbf{Y}_i \boldsymbol{\Omega}_i^{-1} = \mathbf{S}_i & \text{when } n \geq N - 1\end{aligned}$$

$$\mathbf{Y}_i = \mathbf{g}(\mathbf{X}_i) \left(\mathbf{I} - \frac{1}{N} \mathbf{1} \mathbf{1}^T \right) \Big/ \sqrt{N - 1}$$

$$\boldsymbol{\Omega}_i = \mathbf{I} + \mathbf{W}_i \left(\mathbf{I} - \frac{1}{N} \mathbf{1} \mathbf{1}^T \right) \Big/ \sqrt{N - 1}$$

Orthogonal projection $\mathbf{Y}_i \mathbf{A}_i^+ \mathbf{A}_i$

Nonlinear case with $n \geq N - 1$

$$\mathbf{Y}_i \mathbf{A}_i^+ \mathbf{A}_i = \mathbf{Y}_i \left(\mathbf{I} - \frac{1}{N} \mathbf{1} \mathbf{1}^T \right) = \mathbf{Y}_i$$

Nonlinear case with $n < N - 1$

$$\mathbf{Y}_i \mathbf{A}_i^+ \mathbf{A}_i = \mathbf{Y}_i \mathbf{V} \begin{pmatrix} \mathbf{I}_{n \times n} & \mathbf{0} \\ \mathbf{0} & \mathbf{0} \end{pmatrix} \mathbf{V}^T.$$

Linear case

$$\begin{aligned} \mathbf{Y}_i \mathbf{A}_i^+ \mathbf{A}_i &= \mathbf{H} \mathbf{M} \mathbf{X}_i \left(\mathbf{I} - \frac{1}{N} \mathbf{1} \mathbf{1}^T \right) \Big/ \sqrt{N-1} \mathbf{A}_i^+ \mathbf{A}_i \\ &= \mathbf{H} \mathbf{M} \mathbf{A}_i \mathbf{A}_i^+ \mathbf{A}_i \\ &= \mathbf{H} \mathbf{M} \mathbf{A}_i = \mathbf{Y}_i \end{aligned}$$

Equation for \mathbf{W}

Matrix form with $\mathbf{S}_i = \mathbf{Y}_i \boldsymbol{\Omega}_i^{-1}$

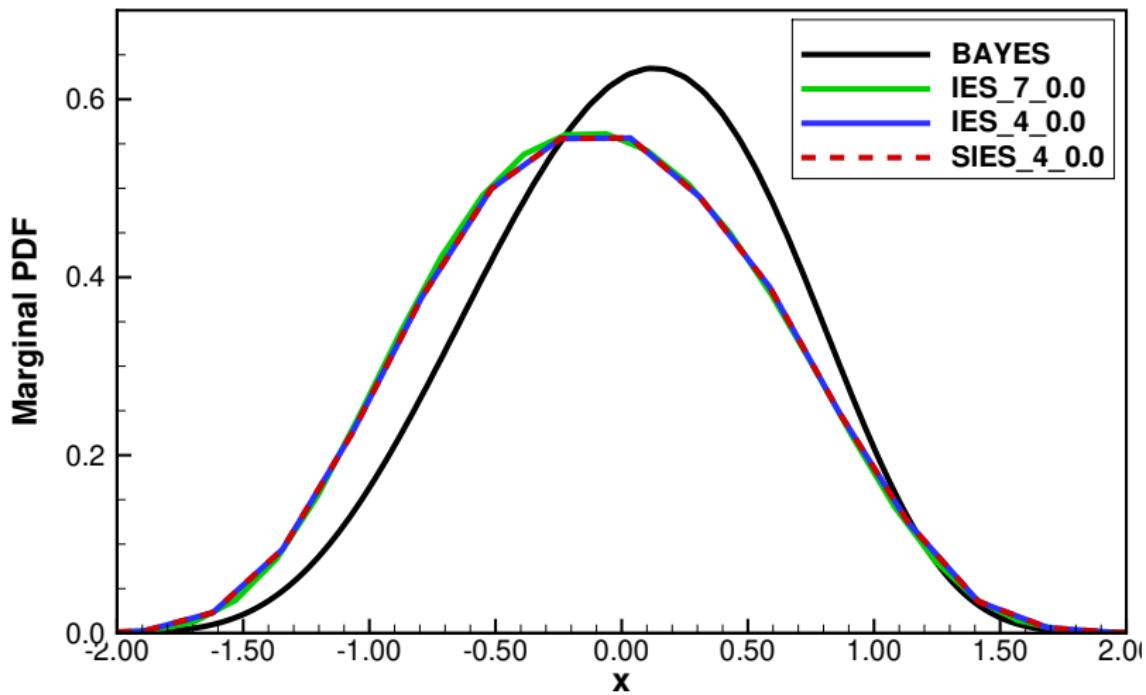
$$\mathbf{W}_{i+1} = \mathbf{W}_i - \gamma \left(\mathbf{W}_i - \mathbf{S}_i^T (\mathbf{S}_i \mathbf{S}_i^T + \mathbf{C}_{dd})^{-1} (\mathbf{S}_i \mathbf{W}_i - \mathbf{D} + \mathbf{g}(\mathbf{X}_i)) \right)$$

IES ensemble subspace algorithm

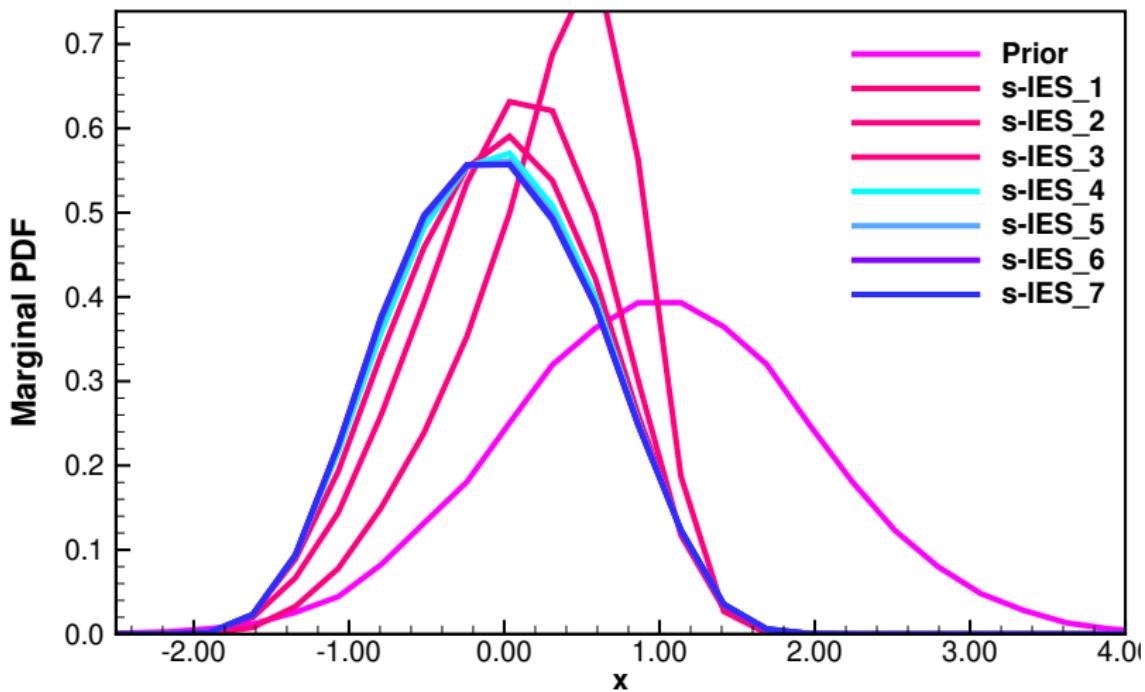
- 1: Inputs: \mathbf{X} , \mathbf{D} , (and \mathbf{C}_{dd})
- 2: $\mathbf{W}_1 = 0$
- 3: **for** $i = 1$, Convergence **do**
- 4: $\mathbf{Y}_i = \mathbf{g}(\mathbf{X}_i)(\mathbf{I} - \frac{1}{N}\mathbf{1}\mathbf{1}^T)/\sqrt{N-1}$
- 5: $\Omega_i = \mathbf{I} + \mathbf{W}_i(\mathbf{I} - \frac{1}{N}\mathbf{1}\mathbf{1}^T)/\sqrt{N-1}$
- 6: $\Omega_i^T \mathbf{S}_i^T = \mathbf{Y}_i^T$ $\mathcal{O}(mN^2)$
- 7: $\mathbf{H}_i = \mathbf{S}_i \mathbf{W}_i + \mathbf{D} - \mathbf{g}(\mathbf{X}_i)$ $\mathcal{O}(mN^2)$
- 8: $\mathbf{W}_{i+1} = \mathbf{W}_i - \gamma \left(\mathbf{W}_i - \mathbf{S}_i^T (\mathbf{S}_i \mathbf{S}_i^T + \mathbf{C}_{dd})^{-1} \mathbf{H}_i \right)$ $\mathcal{O}(mN^2)$
- 9: $\mathbf{X}_{i+1} = \mathbf{X}(\mathbf{I} + \mathbf{W}_{i+1}/\sqrt{N-1})$ $\mathcal{O}(nN^2)$
- 10: **end for**

- ▶ Order $\mathcal{O}(mN^2)$ and $\mathcal{O}(nN^2)$
- ▶ No pseudo inversions of large matrices.

Example nonlinear model



Iterations nonlinear model



Equation for \mathbf{W}

Standard form ($\mathcal{O}(m^3)$)

$$\mathbf{W}_{i+1} = \mathbf{W}_i - \gamma \left(\mathbf{W}_i - \mathbf{S}_i^T (\mathbf{S}_i \mathbf{S}_i^T + \mathbf{C}_{dd})^{-1} \mathbf{H}_i \right)$$

From Woodbury, rewrite as

$$\mathbf{W}_{i+1} = \mathbf{W}_i - \gamma \left\{ \mathbf{W}_i - (\mathbf{S}_i^T \mathbf{C}_{dd}^{-1} \mathbf{S}_i + \mathbf{I}_N)^{-1} \mathbf{S}_i^T \mathbf{C}_{dd}^{-1} \mathbf{H} \right\}$$

For $\mathbf{C}_{dd} = \mathbf{I}_m$ we have ($\mathcal{O}(mN^2)$)

$$\mathbf{W}_{i+1} = \mathbf{W}_i - \gamma \left\{ \mathbf{W}_i - (\mathbf{S}_i^T \mathbf{S}_i + \mathbf{I}_N)^{-1} \mathbf{S}_i^T \mathbf{H} \right\}$$

Subspace inversion (?)

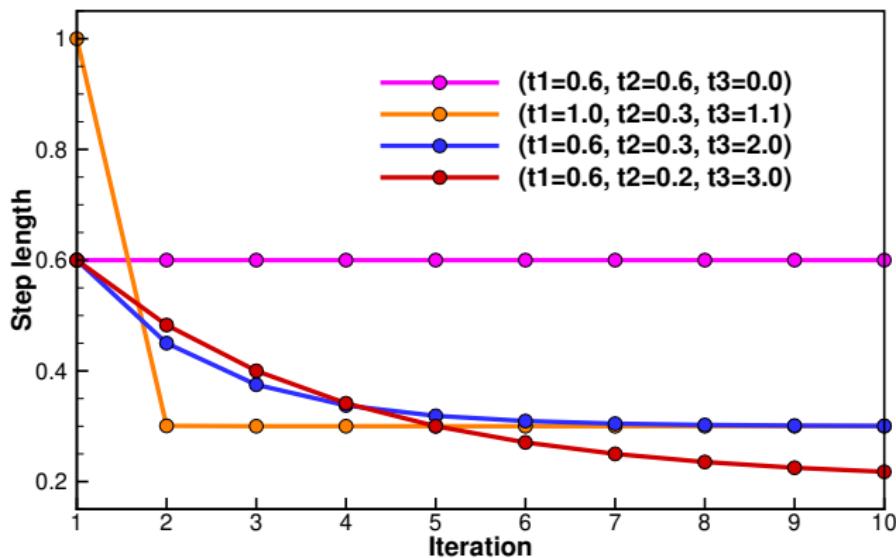
- ▶ Why invert m -dimensional matrix when solving for N coefficients?
- ▶ Do not form $\mathbf{C}_{dd} \approx \mathbf{E}\mathbf{E}^T$ but work directly with \mathbf{E} .

$$\begin{aligned} & (\mathbf{S}\mathbf{S}^T + \mathbf{E}\mathbf{E}^T) \\ & \approx \mathbf{S}\mathbf{S}^T + (\mathbf{S}\mathbf{S}^+)\mathbf{E}\mathbf{E}^T(\mathbf{S}\mathbf{S}^+)^T \\ & = \mathbf{U}\boldsymbol{\Sigma}(\mathbf{I}_N + \boldsymbol{\Sigma}^+\mathbf{U}^T\mathbf{E}\mathbf{E}^T\mathbf{U}(\boldsymbol{\Sigma}^+)^T)\boldsymbol{\Sigma}^T\mathbf{U}^T \\ & = \mathbf{U}\boldsymbol{\Sigma}(\mathbf{I}_N + \mathbf{Z}\boldsymbol{\Lambda}\mathbf{Z}^T)\boldsymbol{\Sigma}^T\mathbf{U}^T \\ & = \mathbf{U}\boldsymbol{\Sigma}\mathbf{Z}(\mathbf{I}_N + \boldsymbol{\Lambda})\mathbf{Z}^T\boldsymbol{\Sigma}^T\mathbf{U}^T. \end{aligned}$$

$$(\mathbf{S}\mathbf{S}^T + \mathbf{E}\mathbf{E}^T)^{-1} \approx \mathbf{U}(\boldsymbol{\Sigma}^+)^T\mathbf{Z}(\mathbf{I}_N + \boldsymbol{\Lambda})^{-1}(\mathbf{U}(\boldsymbol{\Sigma}^+)^T\mathbf{Z})^T$$

- ▶ Cost is $\mathcal{O}(mN^2)$.

Steplength scheme



$$\gamma_i = b + (a - b)2^{(-(i-1)/(c-1))}$$

ERT: <https://github.com/equinor/ert>

ERT

File View Help

Configure Create Plot Export Data Run Workflow Manage Cases Plugins Run Analysis Load results manually Help

Simulation mode: Iterated Ensemble Smoother

Current case: ES_NORMAL

Runpath: poly_out/real_%d/iter_%d

Number of realizations: 100

Number of iterations: 10

Target case format: ES_NORMAL_%d

Analysis Module: IES_ENKF

Active realizations 0-99

Configuration Summary

FORWARD MODEL
poly_eval

PARAMETERS
COEFFS

OBSERVATIONS
POLY_OBS

ERT: <https://github.com/equinor/ert>

ERT

File View Help

Edit variables

Gauss Newton Maximum Steplength

Gauss Newton Minimum Steplength

Gauss Newton Steplength Decline

A good start is max steplength of 0.6, min steplength of 0.3, and decline of 2.5

A steplength of 1.0 and one iteration results in ES update

Inversion algorithm

*0: Exact inversion with diagonal R=I
1: Subspace inversion with exact R
2: Subspace inversion using R=EE'
3: Subspace inversion using E*

Print extensive log for IES

IES Log File

Singular value truncation

Number of singular values

Include AA projection

Any benefit of using the projection is unclear

Close

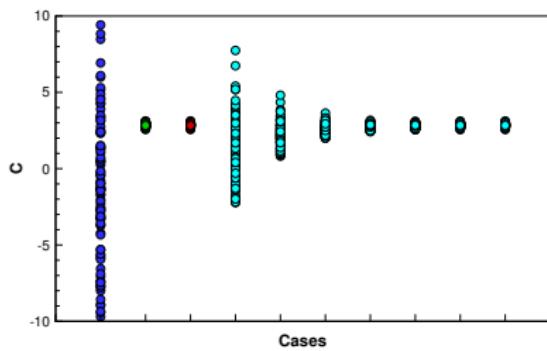
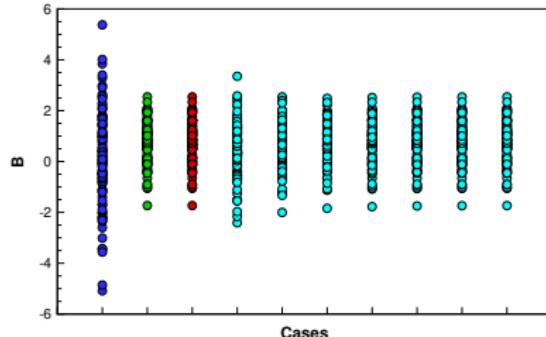
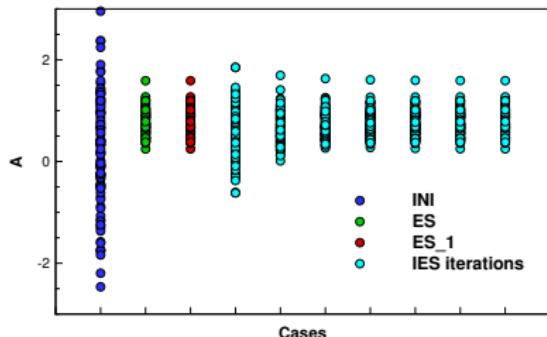
Poly case

Several simple tests are run using a “linear” model

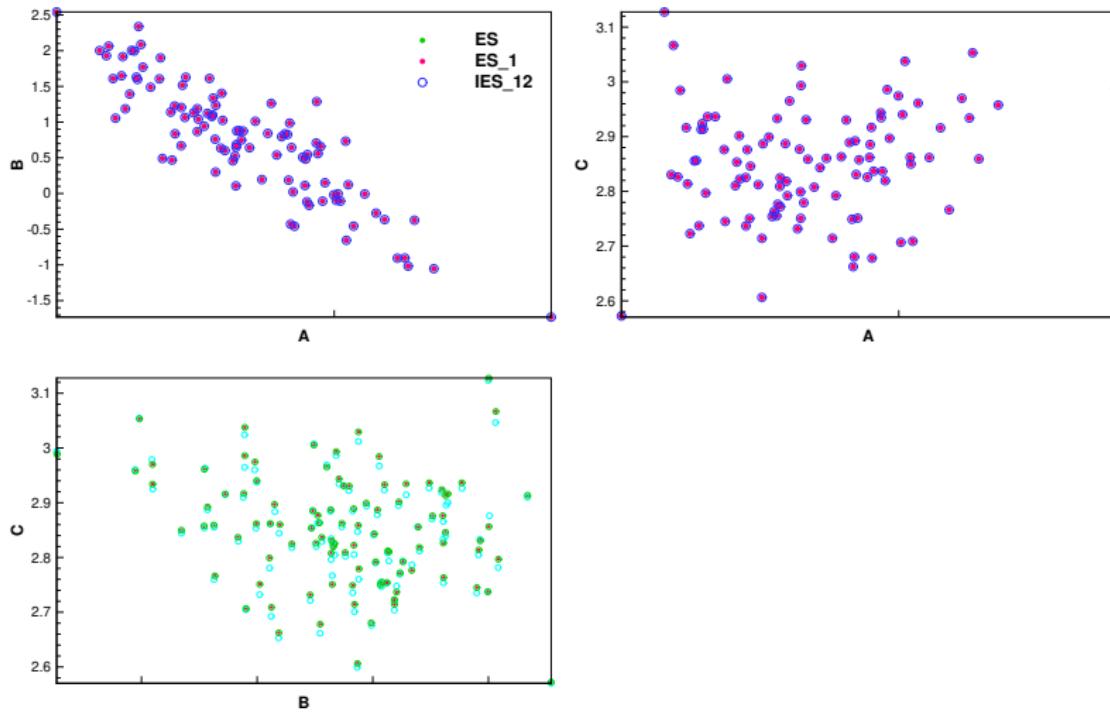
$$y(x) = ax^2 + bx + c \quad (1)$$

- ▶ Coefficients a , b , and c are random Gaussian variables.
- ▶ Measurements (d_1, \dots, d_5) at $x = (0, 2, 4, 6, 8)$.
- ▶ Polynomial curve fitting to the 5 data points.
- ▶ Gauss-linear problem solved exactly by the ES.

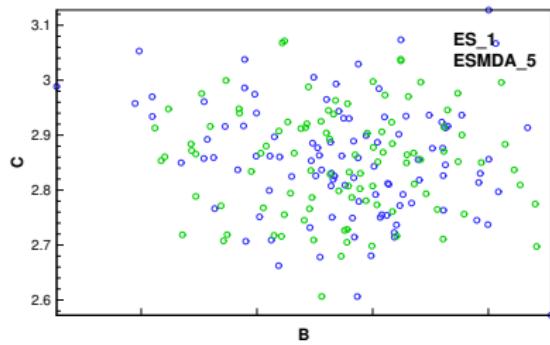
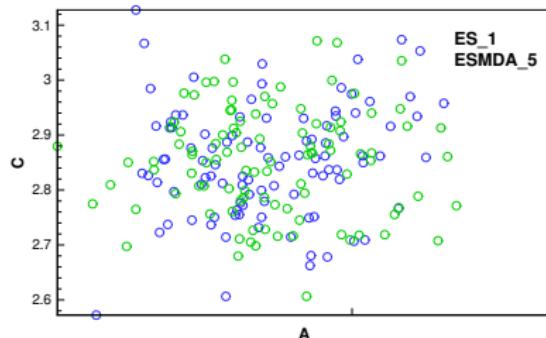
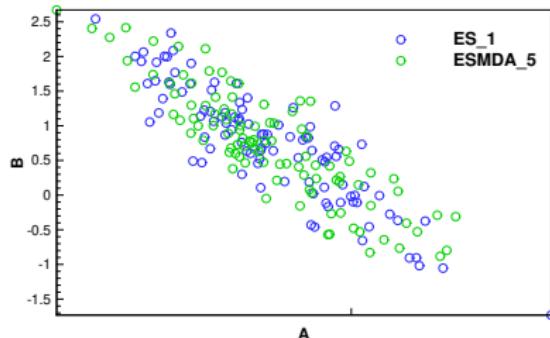
Subspace IES verification



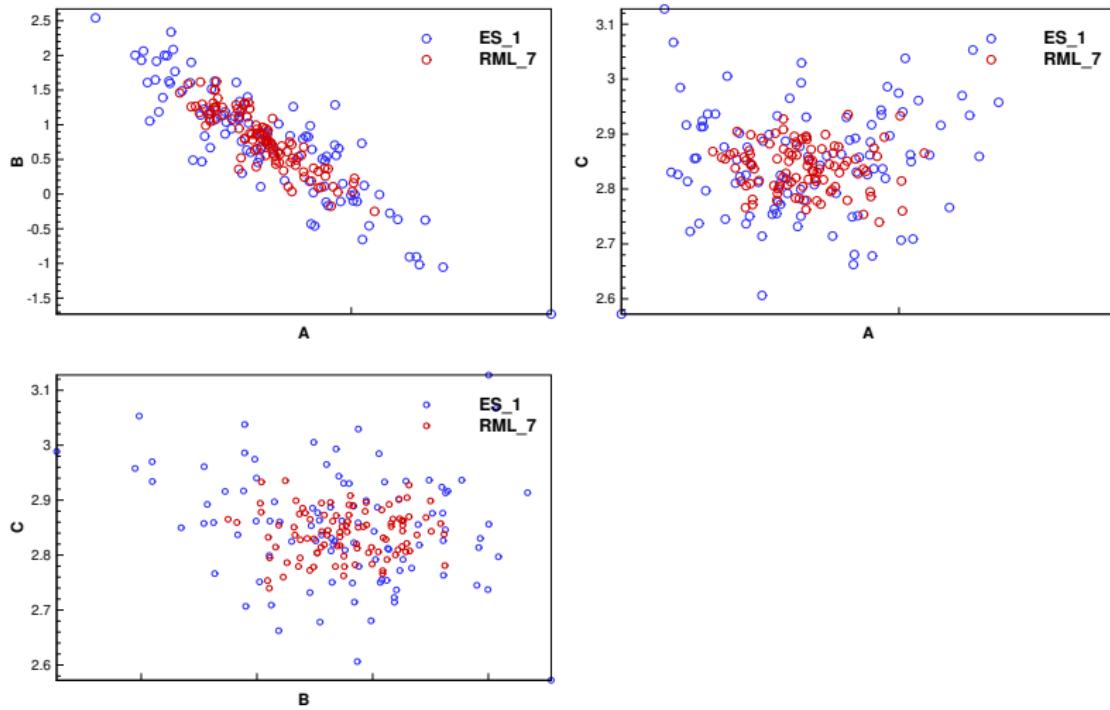
Subspace IES verification



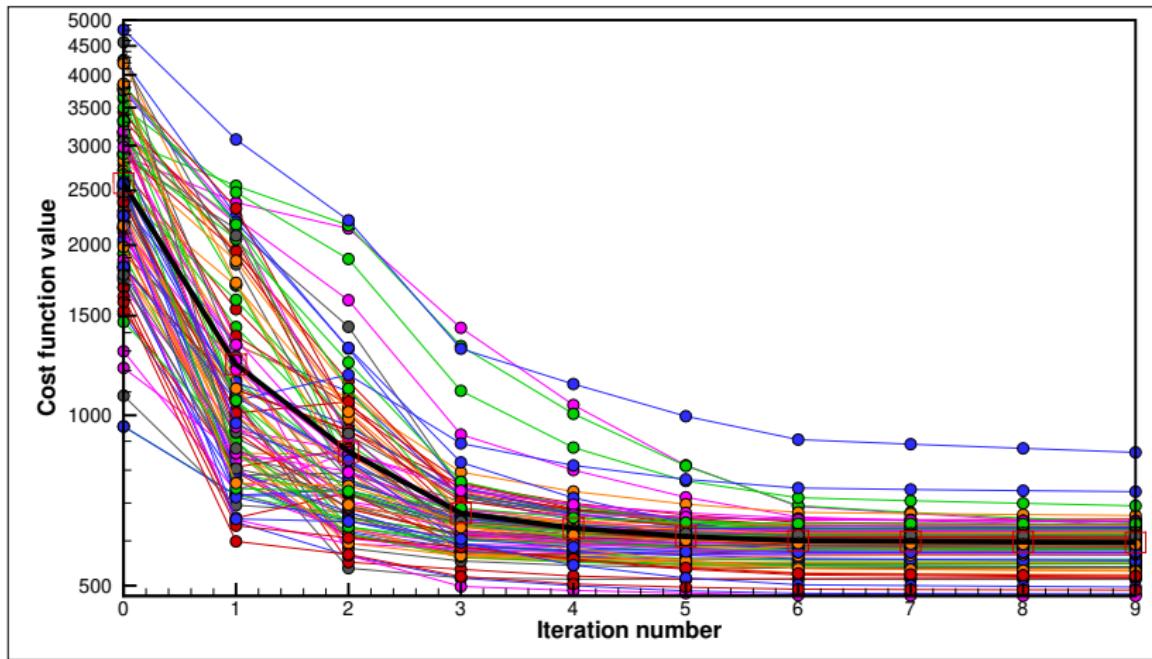
Subspace IES vs. ESMDA



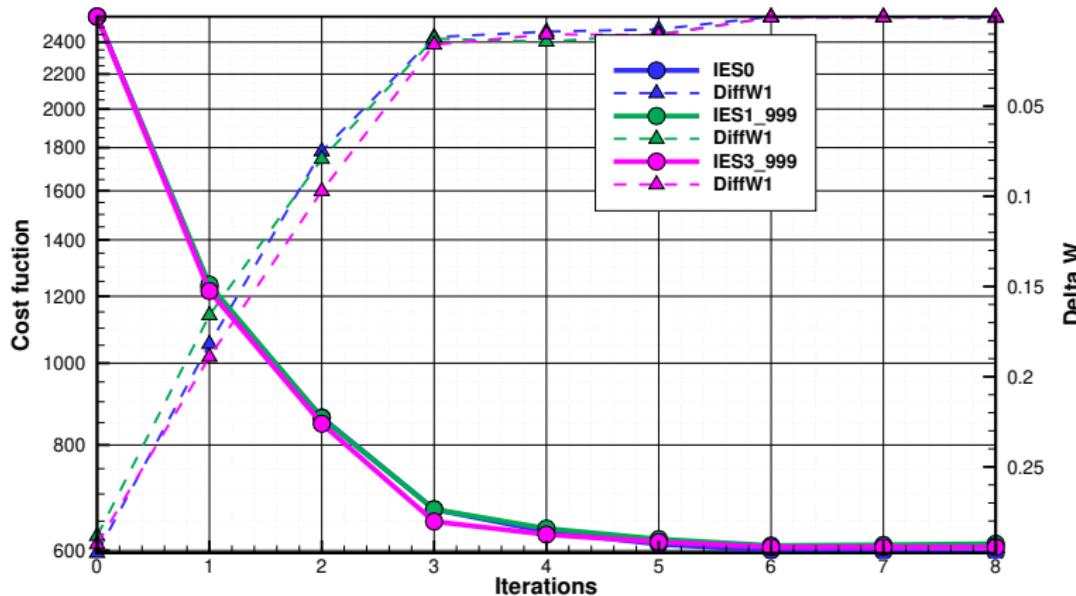
Subspace IES vs. EnRML implementation



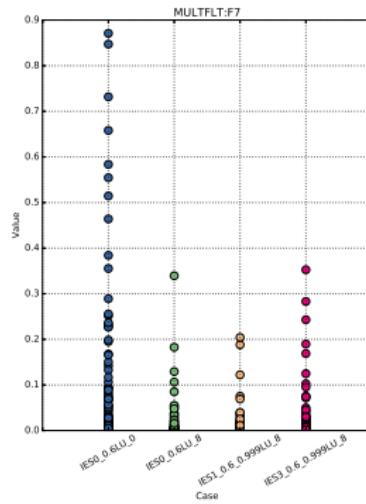
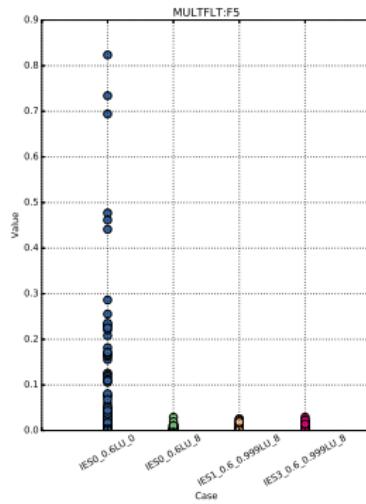
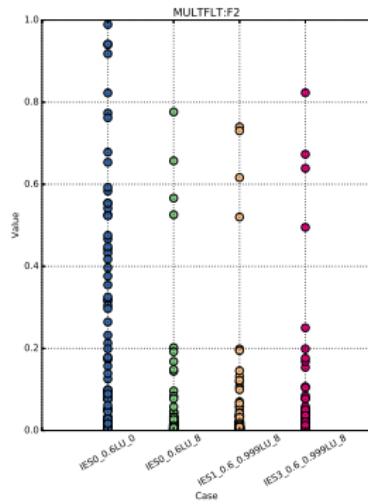
Reek case: Ensemble of cost functions



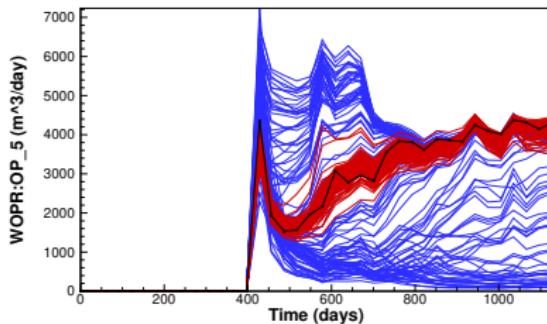
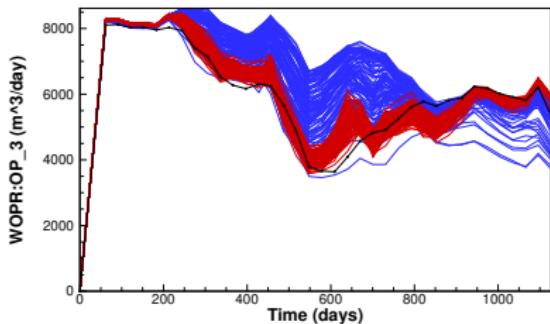
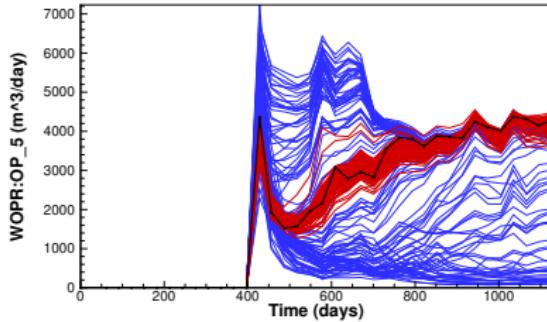
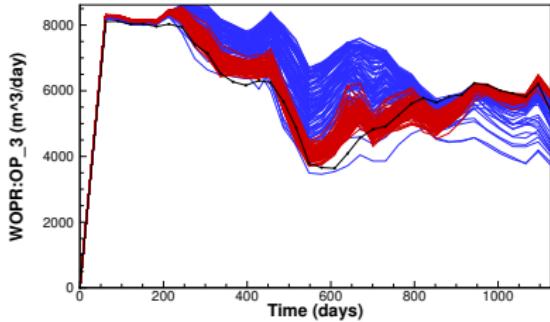
Reek case: Averaged cost function



Reek case: Fault multiplier



Reek case: Oil production



Summary

- ▶ Robust implementation of a robust IES formulation in ERT.
- ▶ IES algorithm formulated for big data and big models.
- ▶ Convergence properties meet requirements for operational use.
- ▶ Pointed out the value of test-based code development.
- ▶ Missing: module for simulating correlated measurement errors.
- ▶ Missing: guidelines for consistent conditioning on production data.

References

<http://digires.no>

<http://digires.no/research-/publications>