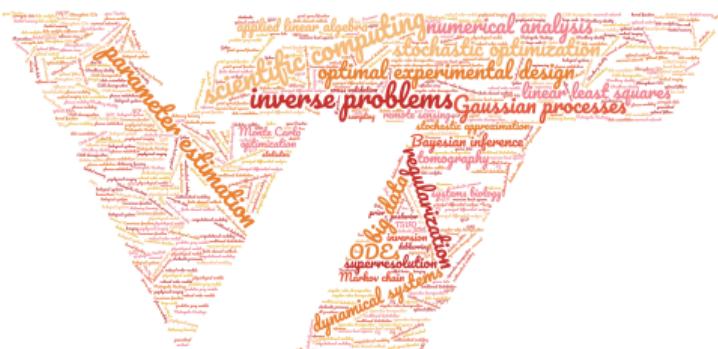


# Challenges in Dynamical Systems Inference: New Approaches for Parameter and Uncertainty Estimation



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Computational Modeling & Data Analytics Division  
Academy of Integrated Science

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January 2020 @ Universität Potsdam

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“All models are wrong, but some are useful, ...

(George Box)

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... and some are dangerous.”

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## Outline

- ① robust ODE/DDE solvers
- ② dynamical systems parameter estimation
- ③ surrogate data
- ④ optimal experimental design

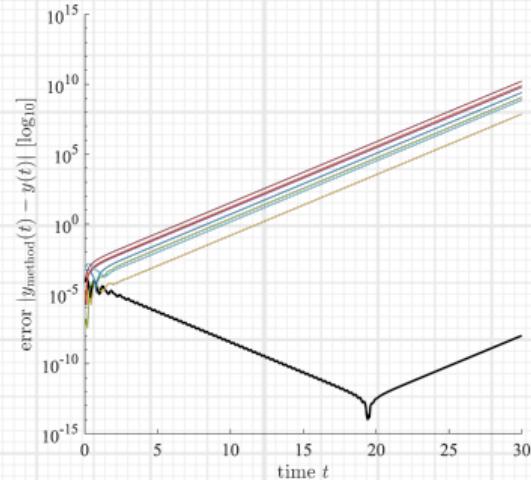
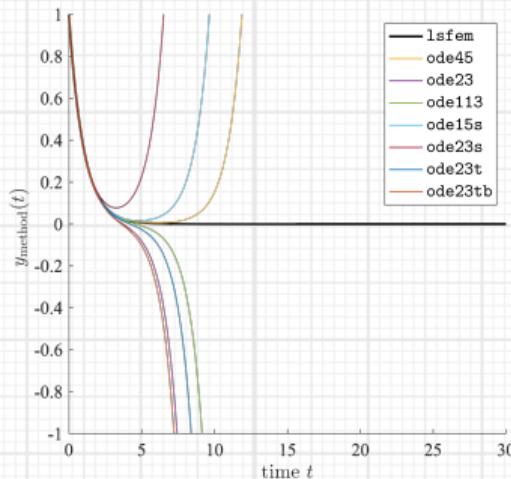
## ① robust lsfem solvers for

- ordinary & delay differential equations (ODE/DDE)
  - initial & boundary value problems (IVP/BVP)
  - differential algebraic equations (DAE)
-

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- ordinary & delay differential equations (ODE/DDE)
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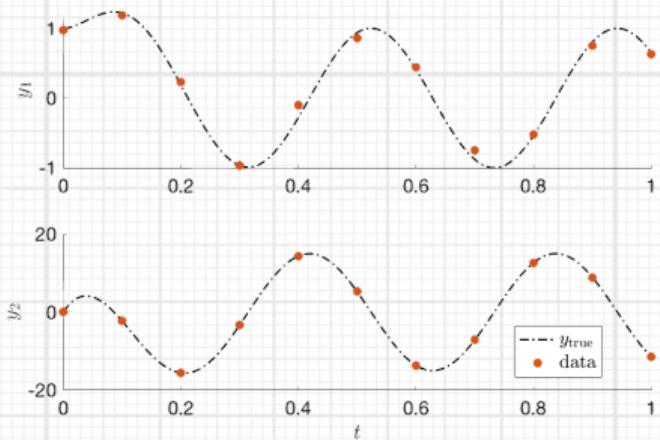
Example  $y' = y - 2e^{-t}$ ,  $y(0) = 1$  with exact solution  $y(t) = e^{-t}$



# ① ODE: Burlisch-Bock parameter estimation problem

$$y' = \begin{bmatrix} 0 & 1 \\ \mu^2 & 0 \end{bmatrix} y - \begin{bmatrix} 0 \\ (\mu^2 + \rho^2) \sin \rho t \end{bmatrix} \quad y(0) = \begin{bmatrix} 1 \\ 0 \end{bmatrix}$$

$t \in [0, 1]$  with  $\mu = \rho = 15$ . Estimate true  $\mu$  for data



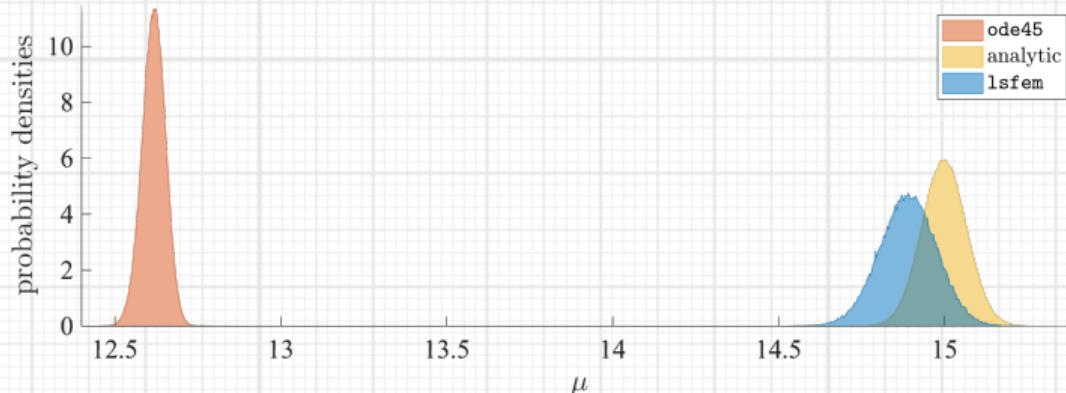
Hans Georg Bock. Recent advances in parameter identification techniques for ode. In: *Numerical treatment of inverse problems in differential and integral equations*. Springer, 1983, pp. 95–121.

R Burlirsch. Die Mehrzielmethode zur numerischen Lösung von nichtlinearen Randwertproblemen und Aufgaben der optimalen Steuerung. In: *Report der Carl-Cranz-Gesellschaft* 251 (1971).

# ① ODE: Burlisch-Bock parameter estimation problem

$$y' = \begin{bmatrix} 0 & 1 \\ \mu^2 & 0 \end{bmatrix} y - \begin{bmatrix} 0 \\ (\mu^2 + \rho^2) \sin \rho t \end{bmatrix} \quad y(0) = \begin{bmatrix} 1 \\ 0 \end{bmatrix}$$

$t \in [0, 1]$  with  $\mu = \rho = 15$ . Results using MCMC method



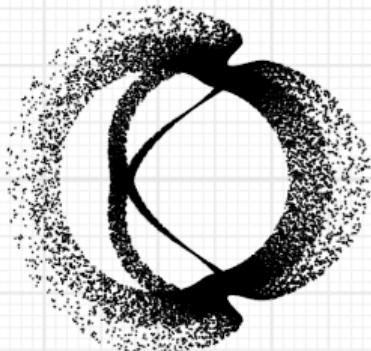
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Bock 1983.

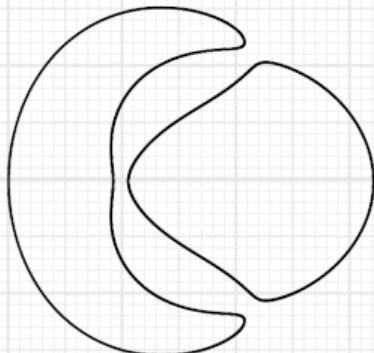
Bulirsch 1971.

# ① Hénon-Heiles Hamiltonian system

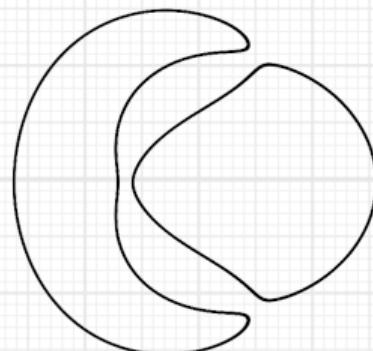
8th order Runge-Kutta



geometric numerical integration



lsfem



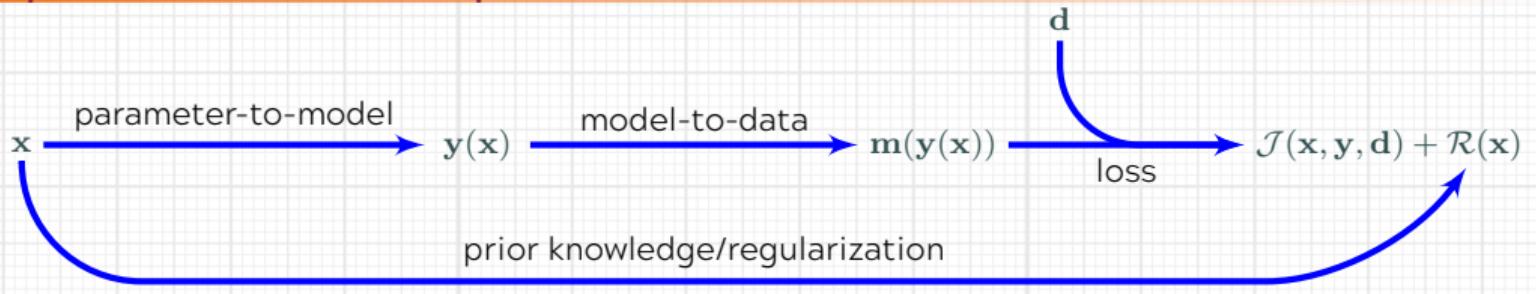
$$\begin{bmatrix} y_1 \\ y_2 \\ y_3 \\ y_4 \end{bmatrix} = \begin{bmatrix} -y_3 - 2\lambda y_3 y_4 \\ -y_4 - \lambda(y_3^2 - y_4^2) \\ y_1 \\ y_2 \end{bmatrix}$$

wrt.

$$\begin{bmatrix} y_1(0) \\ y_2(0) \\ y_3(0) \\ y_4(0) \end{bmatrix} = 0.18 \begin{bmatrix} 1 \\ 1 \\ 1 \\ 1 \end{bmatrix}, \lambda = 1 \text{ and } t \in [0, 100, 000]$$

E Hairer and M Hairer. GniCodes—Matlab programs for geometric numerical integration. In: *Frontiers in Numerical Analysis (Durham, 2002)*. Universitext. Springer, Berlin, 2003, pp. 199–240.

## ② parameter estimation problem



- $\mathbf{x} \in \mathbb{R}^n$  parameter
- $\mathbf{y} : \mathbb{R}^n \rightarrow \mathcal{Y}$  model, e.g., ODE, DDE, PDE
- $\mathbf{m} : \mathcal{Y} \rightarrow \mathbb{R}^m$ , projection onto data, e.g., ODE evaluated at discrete time points
- $\mathbf{d} \in \mathbb{R}^m$
- e.g.,  $\mathcal{J}(\mathbf{x}) = \|\mathbf{m}(\mathbf{y}(\mathbf{x})) - \mathbf{d}\|_2^2$ , where  $\|\cdot\|_2$  Euclidean norm
- $\mathcal{R} : \mathbb{R}^n \rightarrow \mathbb{R}$  regularization/prior knowledge (e.g., sparsity  $\|\cdot\|_1$ )
- Interpretation as MAP estimate in a Bayesian framework

## ② parameter estimation via principal differential analysis (PDA)

$$\hat{\mathbf{x}} = \arg \min_{\mathbf{x}} \|\mathbf{m}(\mathbf{y}(t)) - \mathbf{d}\|_2^2 \quad \text{subject to } \mathbf{y}'(t) = \mathbf{f}(t, \mathbf{y}(t); \mathbf{x}), \quad \mathbf{y}(t_0) = \mathbf{y}_0$$

"derivation":

---

M Chung, J Krueger, and M Pop. Identification of microbiota dynamics using robust parameter estimation methods. In: *Mathematical Biosciences* 294 (2017), pp. 71–84.

Jim O Ramsay. Principal differential analysis: Data reduction by differential operators. In: *Journal of the Royal Statistical Society. Series B (Methodological)* (1996), pp. 495–508.

A.A. Poyton, M.S. Varziri, K.B. McAuley, P.J. McLellan, and J.O. Ramsay. Parameter estimation in continuous-time dynamic models using principal differential analysis. In: *Comput. Chem. Eng.* 30.4 (2006), pp. 698–708.

## ② parameter estimation via principal differential analysis (PDA)

$$(\hat{\mathbf{x}}, \hat{\mathbf{y}}) = \arg \min_{\mathbf{x}, \mathbf{y}} \|\mathbf{m}(\mathbf{y}(t)) - \mathbf{d}\|_2^2 + \lambda \|\mathbf{y}'(t) - \mathbf{f}(t, \mathbf{y}(t); \mathbf{x})\|_{\mathcal{L}^2}^2 \quad \text{subject to } \mathbf{y}(t_0) = \mathbf{y}_0$$

"derivation":

1. relax ODE constraint

---

Chung, Krueger, and Pop 2017.

Ramsay 1996.

Poyton, Varziri, McAuley, McLellan, and Ramsay 2006.

## ② parameter estimation via principal differential analysis (PDA)

$$(\hat{\mathbf{x}}, \hat{\mathbf{q}}) = \arg \min_{\mathbf{x}, \mathbf{q}} \|\mathbf{m}(\mathbf{s}(t; \mathbf{q})) - \mathbf{d}\|_2^2 + \lambda \|\mathbf{s}'(t; \mathbf{q}) - \mathbf{f}(t, \mathbf{s}(t; \mathbf{q}); \mathbf{x})\|_{\mathcal{L}^2}^2, \quad \text{subject to } \mathbf{s}(t_0) = \mathbf{y}_0$$

"derivation":

1. relax ODE constraint
2. restrict to parameterized finite function space

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"derivation":

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2. restrict to parameterized finite function space
3. discretize  $\mathbf{T} = [T_1, \dots, T_M]^\top$

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"derivation":

1. relax ODE constraint
2. restrict to parameterized finite function space
3. discretize  $\mathbf{T} = [T_1, \dots, T_M]^\top$

advantages:

- simultaneous parameter and approximate ODE solve
- robustness in parameter estimates

---

Chung, Krueger, and Pop 2017.

Ramsay 1996.

Poyton, Varziri, McAuley, McLellan, and Ramsay 2006.

## ② generalized Lotka-Volterra simulation study

Consider the 4 state Lotka-Volterra system

$$\mathbf{y}' = \text{diag}(\mathbf{y})(\mathbf{r} + \mathbf{A}\mathbf{y}), \quad \mathbf{y}(0) = \mathbf{y}_0$$

with

$$\mathbf{r} = \begin{bmatrix} 2 \\ 1 \\ 0 \\ -3 \end{bmatrix}, \quad \mathbf{A} = \begin{bmatrix} 0 & -0.6 & 0 & -0.2 \\ 0.6 & 0 & -0.6 & -0.2 \\ 0 & 0.6 & 0 & -0.2 \\ 0.2 & 0.2 & 0.2 & 0 \end{bmatrix}, \quad \mathbf{y}_0 = \begin{bmatrix} 5 \\ 4 \\ 3 \\ 2 \end{bmatrix}.$$

Goal: Estimating  $\mathbf{x} = [\mathbf{r}; \text{vec}(\mathbf{A}); \mathbf{y}_0]$  give data.

## ② data recovery

Average relative error:

$$e_r = \frac{1}{80} \sum_{j=1}^{80} \left| \frac{m_j(\mathbf{y}) - d_j}{d_j} \right|$$

Study 1 (0% noise):

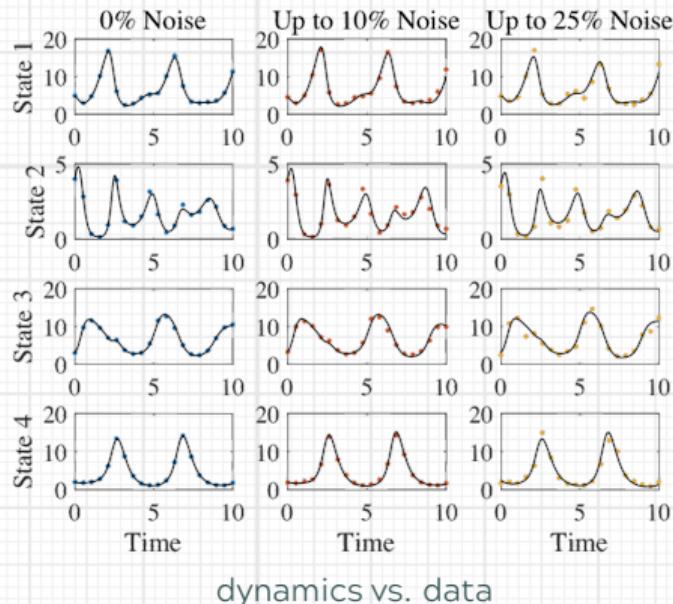
$$e_r \approx 0.0331$$

Study 2 (up to 10% noise):

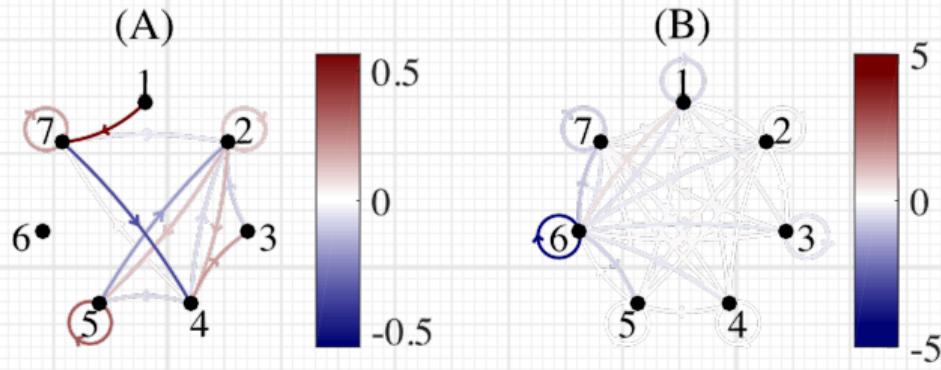
$$e_r \approx 0.0926$$

Study 3 (up to 25% noise):

$$e_r \approx 0.1511$$



## ② intestinal microbiota (interaction matrix and comparison)



PDA (A) versus published interaction matrix (B)

- |                              |                                 |
|------------------------------|---------------------------------|
| 1. Blautia                   | 5. Unclassified Lachnospiraceae |
| 2. Barnesiella               | 6. Coprobacillus                |
| 3. Unclassified Mollicutes   | 7. Other                        |
| 4. Undefined Lachnospiraceae |                                 |

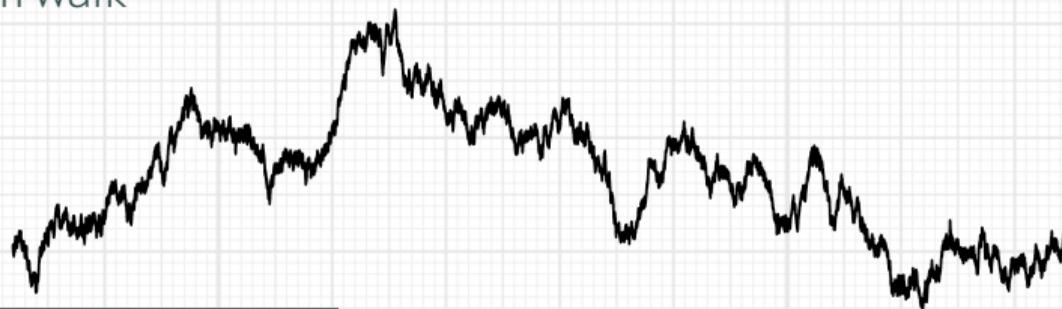
C G Buffie, I Jarchum, et al. Profound alterations of intestinal microbiota following a single dose of clindamycin results in sustained susceptibility to *Clostridium difficile*-induced colitis. In: *Infect. Immun.* 80.1 (2012), pp. 62–73; R R Stein, V Bucci, et al. Ecological modeling from time-series inference: insight into dynamics and stability of intestinal microbiota. In: *PLoS Comput. Biol.* 9.12 (2013), e1003388; C G Buffie, V Bucci, et al. Precision microbiome reconstitution restores bile acid mediated resistance to *Clostridium difficile*. In: *Nature* 517.7533 (2015), pp. 205–208.

### ③ Gaussian processes

A **Gaussian process** is a collection of random variables  $g(t)$ ; any finite number  $\{g(t_i)\}_{i=1}^m$  of which have a joint Gaussian distribution, i.e., for finite  $\mathbf{t} = [t_1, \dots, t_m]^\top \in \mathbb{R}^m$ , the joint distribution is Gaussian,

$$g(\mathbf{t}) \sim \mathcal{N}([\mu(t_i)]_{i=1}^m, [\kappa(t_i, t_j)]_{i,j=1}^m)$$

e.g., random walk



### ③ Gaussian process prior

Kernel  $\kappa(t, \tilde{t})$  is real, symmetric, non-negative, integrable function

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squared exponential kernel  $\kappa(t, \tilde{t}) = \tau^2 \exp\left(-\frac{1}{2\ell^2} \|t - \tilde{t}\|_2^2\right)$  ( $\tau = 1$ )

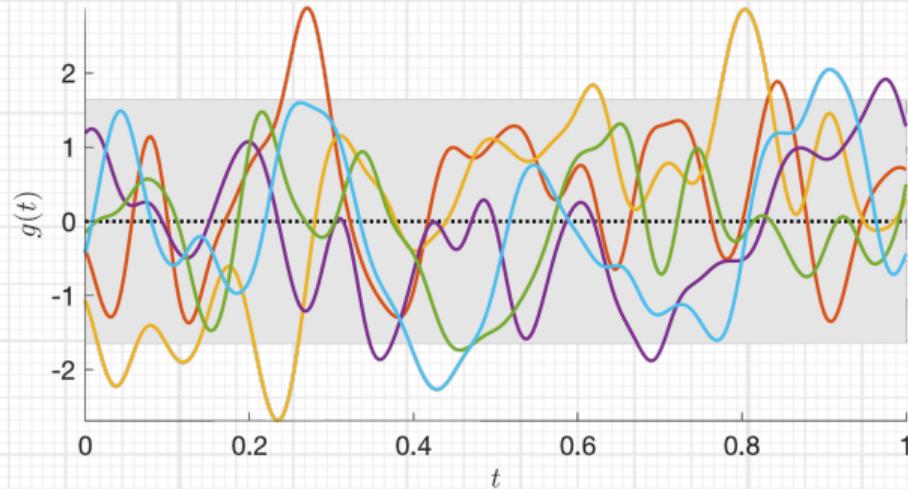


$$\ell^2 = 0.01 \text{ (length scale parameter)}$$

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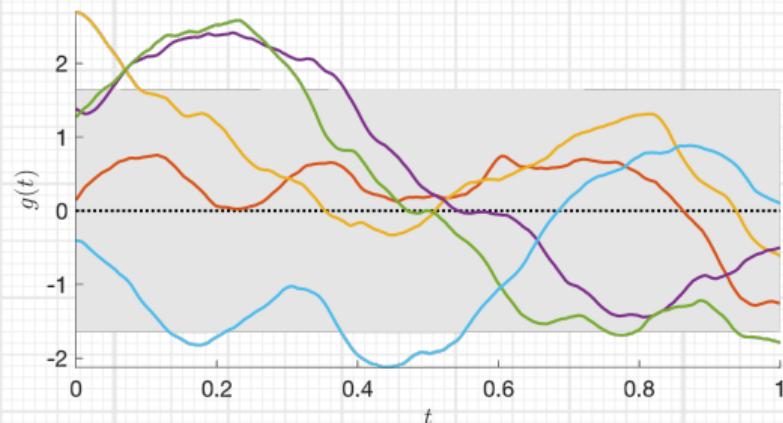


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### ③ Gaussian process prior

Kernel  $\kappa(t, \tilde{t})$  is real, symmetric, non-negative, integrable function

Matérn kernel  $\kappa(t, \tilde{t}) = \frac{2^{1-\nu}}{\Gamma(\nu)} \left( \frac{\sqrt{2\nu|t-\tilde{t}|}}{\ell} \right)^\nu K_\nu \left( \frac{\sqrt{2\nu|t-\tilde{t}|}}{\ell} \right)$

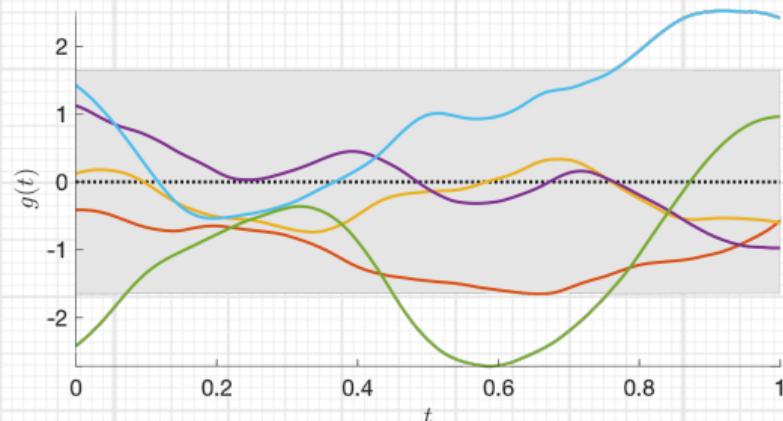


$\Gamma$  Gamma function,  $K_\nu$  modified Bessel function,  $\nu = 1/2$

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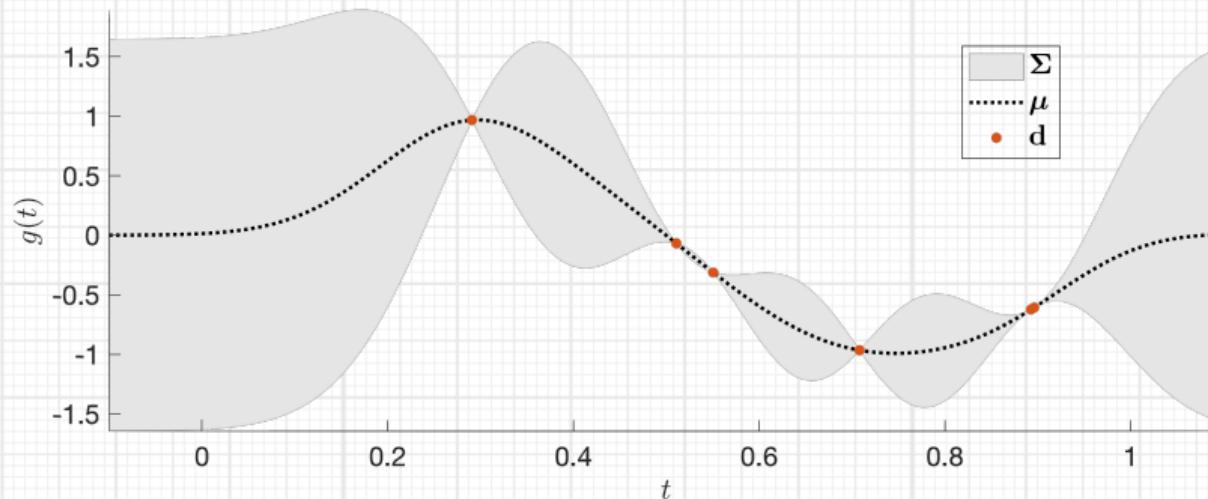
$\Gamma$  Gamma function,  $K_\nu$  modified Bessel function,  $\nu = 3/2$

### ③ conditional distribution/prediction

joint distribution at  $t$  and  $\mathbf{T}$ :  $\begin{bmatrix} \mathbf{d} \\ \mathbf{g} \end{bmatrix} \sim \mathcal{N} \left( \begin{bmatrix} \mathbf{0}_m \\ \mathbf{0}_M \end{bmatrix}, \begin{bmatrix} \Sigma_t & \Sigma_{tT} \\ \Sigma_{Tt} & \Sigma_T \end{bmatrix} \right)$

conditional (predictive) distribution

$$(\mathbf{g}|\mathbf{d}) \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma}), \quad \text{with } \boldsymbol{\mu} = \Sigma_{Tt} \Sigma_t^{-1} \mathbf{d} \quad \text{and } \boldsymbol{\Sigma} = \Sigma_T - \Sigma_{Tt} \Sigma_t^{-1} \Sigma_{tT}$$



### ③ surrogate problem

For predictive Gaussian process  $\mathbf{g} \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$  for given data  $(\mathbf{t}, \mathbf{d})$  and model  $\mathbf{y}$  solve

$$\hat{\mathbf{x}} = \arg \min_{\mathbf{x}} \|\mathbf{m}(\mathbf{y}(t, \mathbf{x})) - \mathbf{g}\|_{\boldsymbol{\Sigma}^{-1}}^2 + \mathcal{R}(\mathbf{x})$$

**Algorithm:** sampled GP weighted least squares

**input:** model  $\mathbf{y}$ , projection  $\mathbf{m}$ ,  $\boldsymbol{\mu}$ ,  $\boldsymbol{\Sigma}$ , initial guess  $\mathbf{x}_0$

1: **parallel for**  $j = 1$  to  $J$  **do**

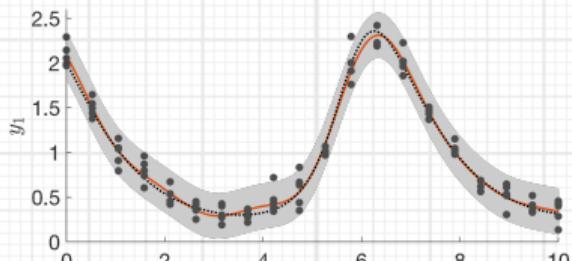
2:     **sample**  $\mathbf{g}_j$  **from Gaussian process**  $\mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

3:     **solve**  $\hat{\mathbf{x}}_j = \arg \min_{\mathbf{x}} \|\mathbf{m}(\mathbf{y}(t, \mathbf{x})) - \mathbf{g}_j\|_{\boldsymbol{\Sigma}^{-1}}^2 + \mathcal{R}(\mathbf{x})$

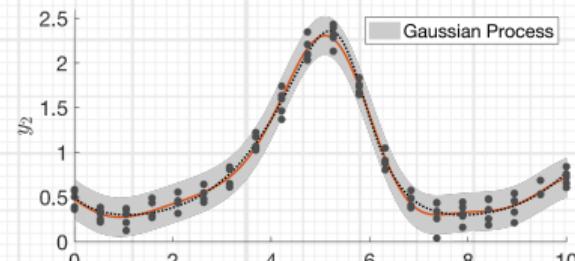
4: **end parallel for**

**output:**  $\{\hat{\mathbf{x}}_j\}_{j=1}^J$

### ③ motivating example: Lotka-Volterra & model

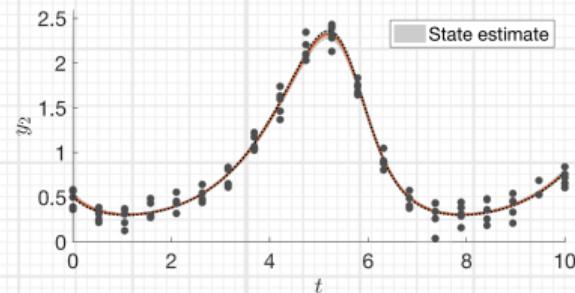
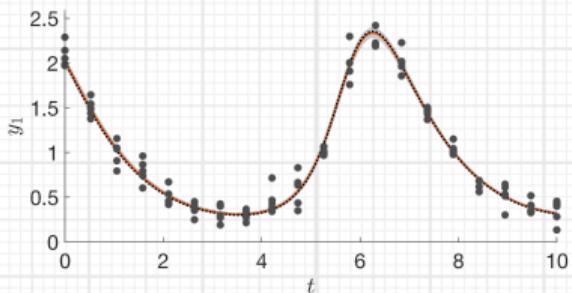


$$y'_1 = -y_1 + x_1 y_1 y_2$$



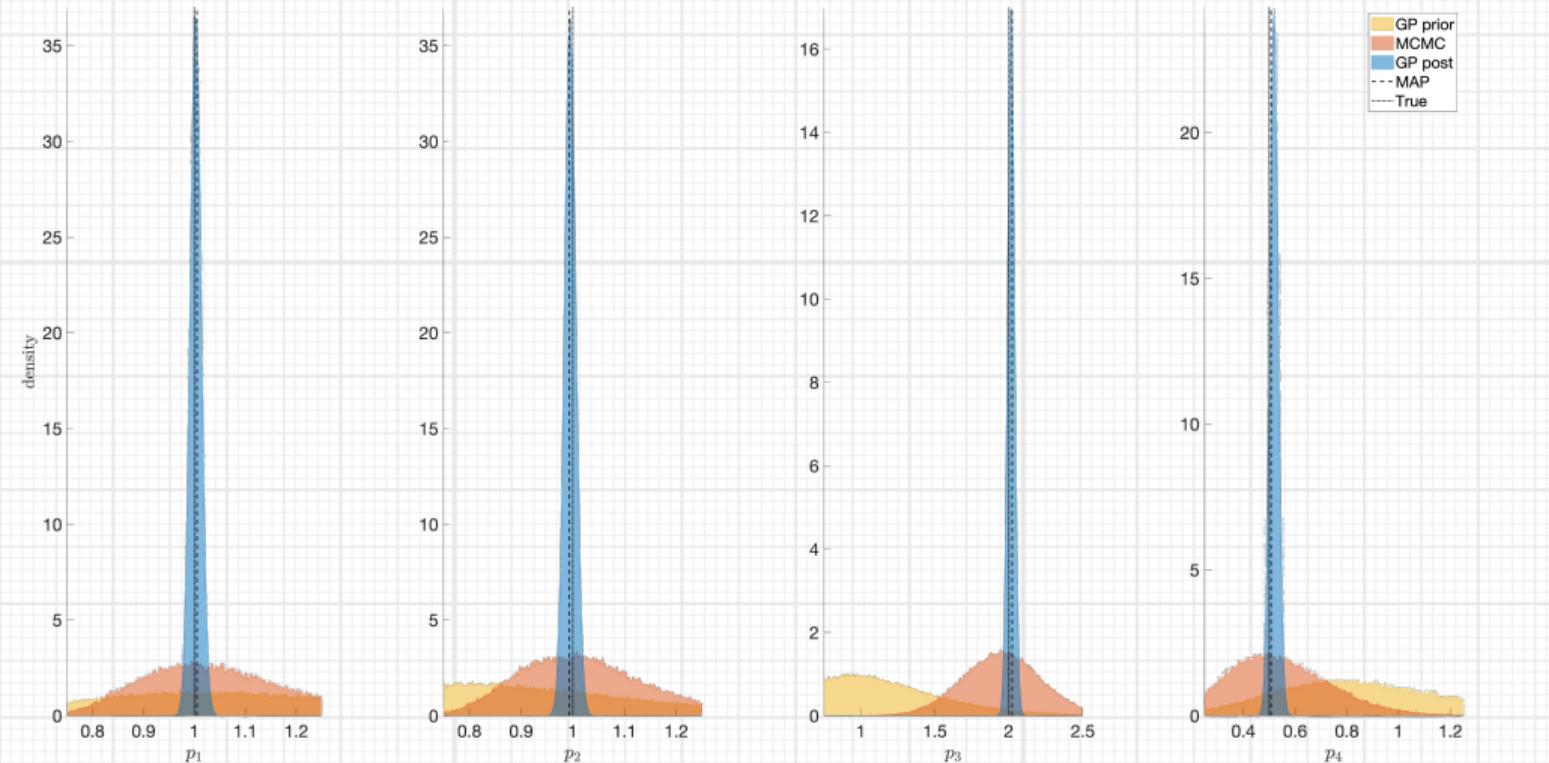
$$y'_2 = y_2 - x_2 y_1 y_2$$

unknown parameter  $\mathbf{x} = [x_1, x_2, y_1(0), y_2(0)]^\top$



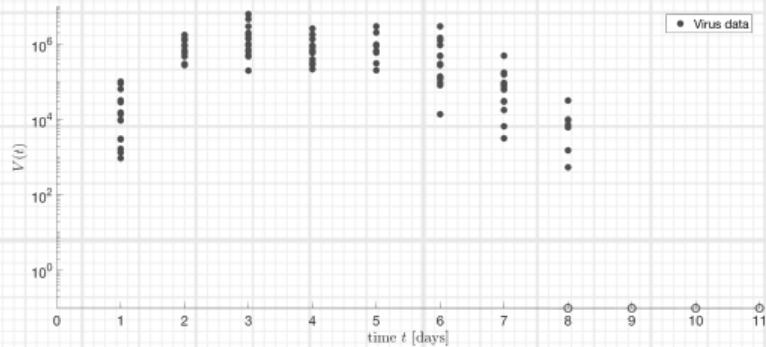
Matthias Chung, Mickaël Binois, et al. Parameter and uncertainty estimation for dynamical systems using surrogate stochastic processes. In: *SIAM Journal on Scientific Computing* 41.4 (2019), A2212–A2238.

### ③ Lotka-Volterra UQ



marginal 1D densities with GP and with MCMC

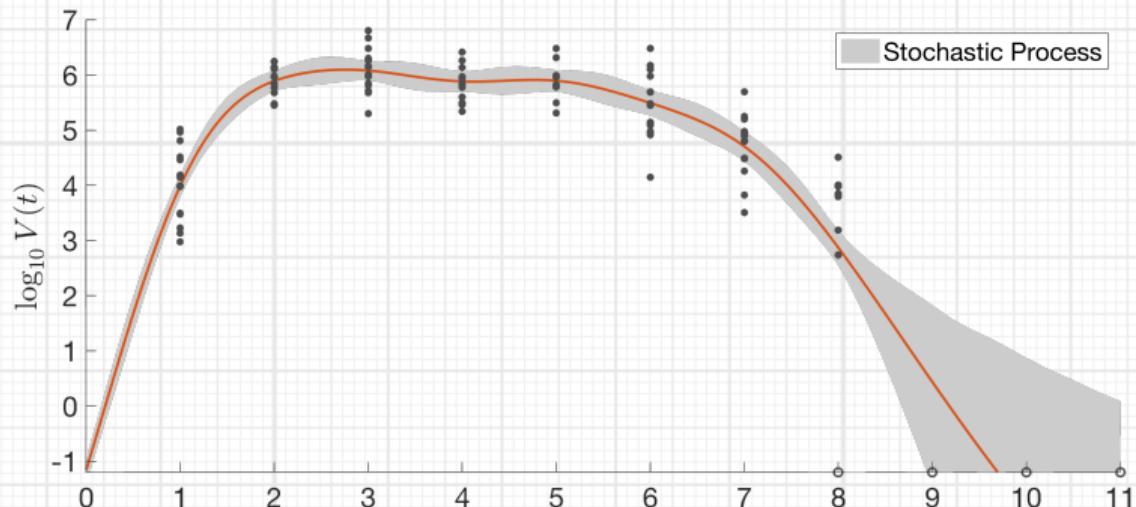
### ③ motivating example: influenza virus data & model



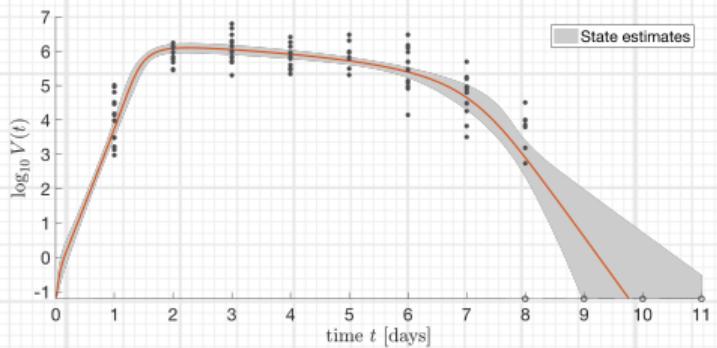
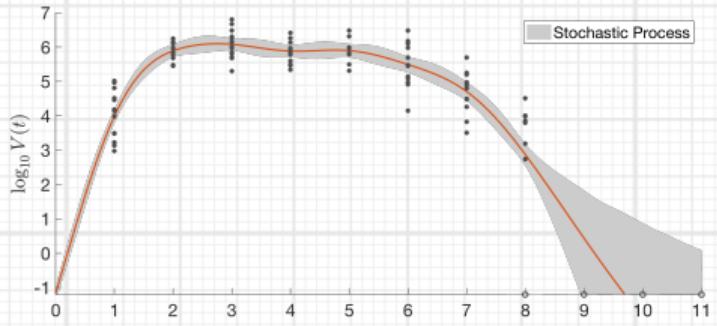
$$\begin{aligned} T' &= -\beta TV \\ I_1' &= \beta TV - \kappa I_1 \\ I_2' &= \kappa I_1 - \frac{\delta I_2}{K_d + I_2} \\ V' &= \rho I_2 - cV \end{aligned}$$

- $T$  target cells
- $I_1$  and  $I_2$  infected cells
- $V$  virus
- $\beta$  "contact rate"
- $\kappa$  rate (eclipse)
- $\rho$  virus production rate
- $c$  clearance rate
- $\delta$  density dependent clearance
- $K_d$  half-saturation constant

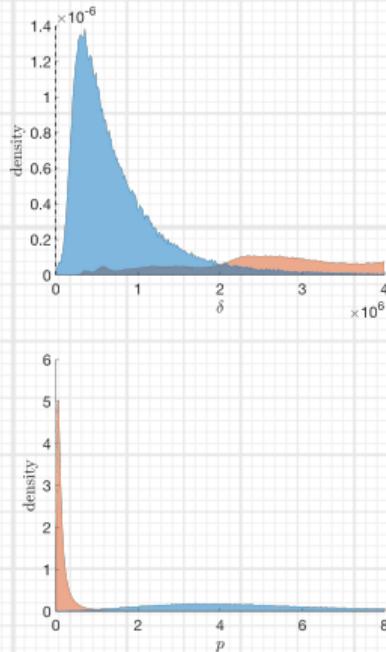
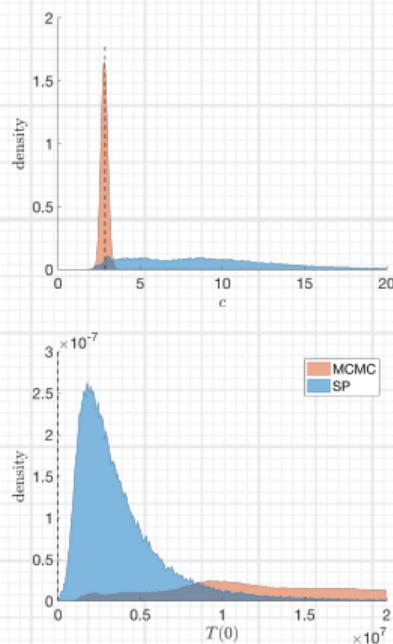
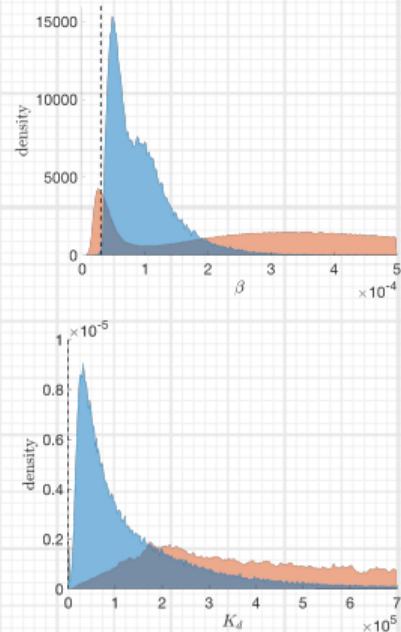
### ③ influenza virus GP



### ③ influenza virus GP



### ③ influenza virus GP



marginal 1D densities with **GP** and with **MCMC**

## ④ optimal experimental design framework

$$\hat{\mathbf{x}} = \arg \min_{\mathbf{x}} \underbrace{\frac{1}{2} \| \mathbf{y} - \mathbf{x} - \mathbf{d} \|_{\Gamma_{\epsilon}^{-1}}^2}_{\mathcal{J}(\mathbf{x})} + \underbrace{\frac{1}{2} \| \mathbf{Lx} \|_2^2}_{\mathcal{R}(\mathbf{x})}$$

subject to  $\mathbf{C}_e \mathbf{x} - \mathbf{c}_e = \mathbf{0}$  and  $\mathbf{C}_i \mathbf{x} - \mathbf{c}_i \geq \mathbf{0}$ ,

- $\mathbf{C}_e, \mathbf{C}_i, \mathbf{c}_e, \mathbf{c}_i$  constraints

---

E Haber, L Horesh, and L Tenorio. Numerical methods for experimental design of large-scale linear ill-posed inverse problems. In: *Inverse Problems* 24.5 (2008), p. 055012.

M Chung and E Haber. Experimental Design for Biological Systems. In: *SIAM Journal on Control and Optimization* 50.1 (2012), pp. 471–489.

## ④ optimal experimental design framework

$$\hat{\mathbf{x}}(\mathbf{p}) = \arg \min_{\mathbf{x}} \underbrace{\frac{1}{2} \|\mathbf{y}(\mathbf{p}, \mathbf{x}) - \mathbf{d}(\mathbf{p})\|_{\Gamma_{\epsilon}^{-1}(\mathbf{p})}^2}_{\mathcal{J}(\mathbf{x})} + \underbrace{\frac{1}{2} \|\mathbf{L}\mathbf{x}\|_2^2}_{\mathcal{R}(\mathbf{x})}$$

subject to  $\mathbf{C}_e \mathbf{x} - \mathbf{c}_e = \mathbf{0}$  and  $\mathbf{C}_i \mathbf{x} - \mathbf{c}_i \geq \mathbf{0}$ ,

- $\mathbf{C}_e, \mathbf{C}_i, \mathbf{c}_e, \mathbf{c}_i$  constraints
- $\mathbf{p} \in \Omega$  design parameter ( $\Omega$  set of feasible design parameter)

## ④ optimal experimental design framework

Bayes risk approach (average design)

$$\min_{\mathbf{p} \in \Omega} \quad \mathcal{J}(\mathbf{p}) = \frac{1}{2} \mathbb{E} \| \widehat{\mathbf{x}}(\mathbf{p}) - \mathbf{x}_{\text{true}} \|_2^2 + \mathcal{R}_{\mathbf{p}}(\mathbf{p}) \quad \text{subject to}$$

$$\widehat{\mathbf{x}}(\mathbf{p}) = \arg \min_{\mathbf{x}} \underbrace{\frac{1}{2} \| \mathbf{y}(\mathbf{p}, \mathbf{x}) - \mathbf{d}(\mathbf{p}) \|_{\Gamma_{\epsilon}^{-1}(\mathbf{p})}^2}_{\mathcal{J}(\mathbf{x})} + \underbrace{\frac{1}{2} \| \mathbf{L}\mathbf{x} \|_2^2}_{\mathcal{R}(\mathbf{x})}$$

subject to  $\mathbf{C}_e \mathbf{x} - \mathbf{c}_e = \mathbf{0}$  and  $\mathbf{C}_i \mathbf{x} - \mathbf{c}_i \geq \mathbf{0}$ ,

- $\mathbf{C}_e, \mathbf{C}_i, \mathbf{c}_e, \mathbf{c}_i$  constraints
- $\mathbf{p} \in \Omega$  design parameter ( $\Omega$  set of feasible design parameter)
- $\mathbb{E}$  expected value (sampling of expected value via training data  $\mathbf{x}_{\text{true}}^j$ , with  $j = 1, \dots, M$ )
- $\mathcal{R}_{\mathbf{p}}$  regularization (cost) on design parameter  $\mathbf{p}$

Haber, Horesh, and Tenorio 2008.

Chung and Haber 2012.

## ④ optimal experimental design framework

Bayes risk approach (average design)

$$\min_{\mathbf{p} \in \Omega} \quad \mathcal{J}(\mathbf{p}) = \frac{1}{2} \mathbb{E} \|\hat{\mathbf{x}}(\mathbf{p}) - \mathbf{x}_{\text{true}}\|_2^2 + \beta \|\mathbf{p}\|_1 \quad \text{subject to}$$

$$\hat{\mathbf{x}}(\mathbf{p}) = \arg \min_{\mathbf{x}} \underbrace{\frac{1}{2} \|\mathbf{y}(\mathbf{p}, \mathbf{x}) - \mathbf{d}(\mathbf{p})\|_{\Gamma_{\epsilon}^{-1}(\mathbf{p})}^2}_{\mathcal{J}(\mathbf{x})} + \underbrace{\frac{1}{2} \|\mathbf{L}\mathbf{x}\|_2^2}_{\mathcal{R}(\mathbf{x})}$$

subject to  $\mathbf{C}_e \mathbf{x} - \mathbf{c}_e = \mathbf{0}$  and  $\mathbf{C}_i \mathbf{x} - \mathbf{c}_i \geq \mathbf{0}$ ,

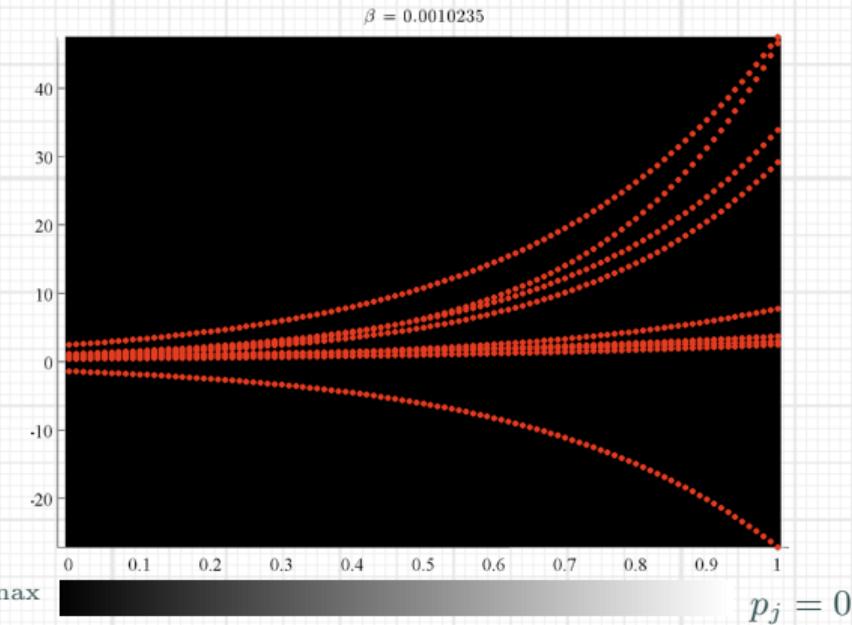
- $\mathbf{C}_e, \mathbf{C}_i, \mathbf{c}_e, \mathbf{c}_i$  constraints
- $\mathbf{p} \in \Omega$  design parameter ( $\Omega$  set of feasible design parameter)
- $\mathbb{E}$  expected value (sampling of expected value via training data  $\mathbf{x}_{\text{true}}^j$ , with  $j = 1, \dots, M$ )
- $\mathcal{R}_{\mathbf{p}}$  regularization (cost) on design parameter  $\mathbf{p}$

Haber, Horesh, and Tenorio 2008.

Chung and Haber 2012.

## ④ Exponential growth

$$y' = xy \quad \text{unknown parameter } x$$



samples = 10, discretization 101, no noise, and  $\mathbf{L} = 0.1\mathbf{I}_n$

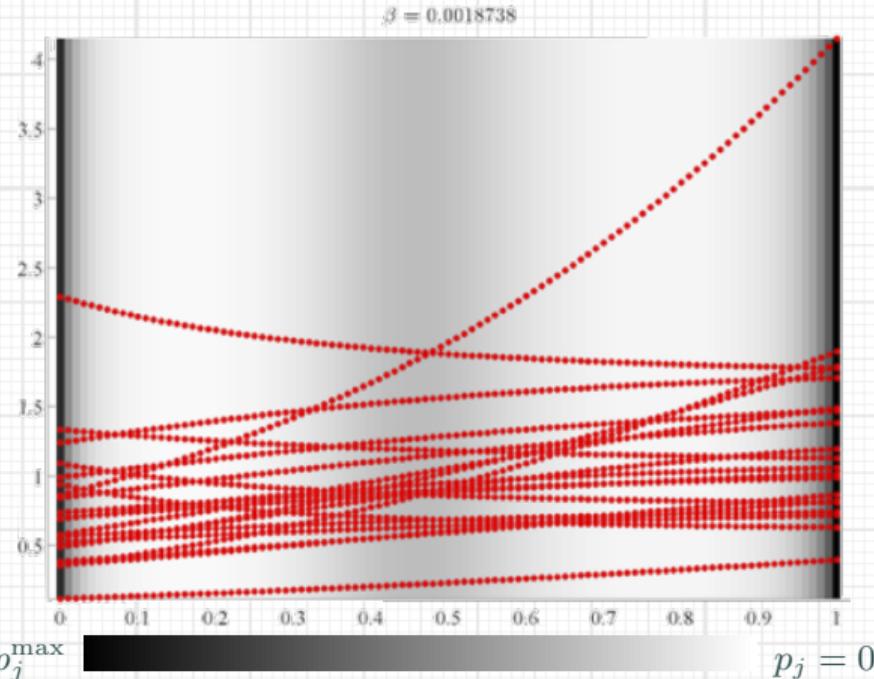
## ④ Exponential growth

$$y' = xy \quad \text{unknown parameter } x$$



## ④ Logistic growth

$$y' = x_1 y - x_2 y^2 \quad \text{unknown parameter } \mathbf{x} = [x_1, x_2]^\top$$



samples = 20, discretization 101, no noise, and  $\mathbf{L} = 0.1\mathbf{I}_n$

## ④ Logistic growth

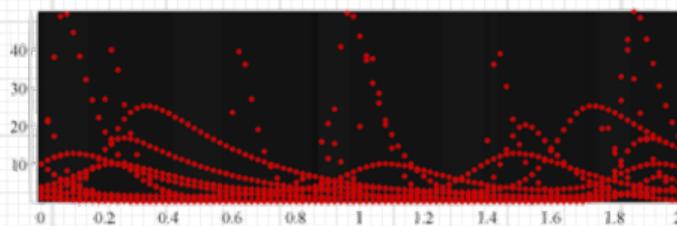
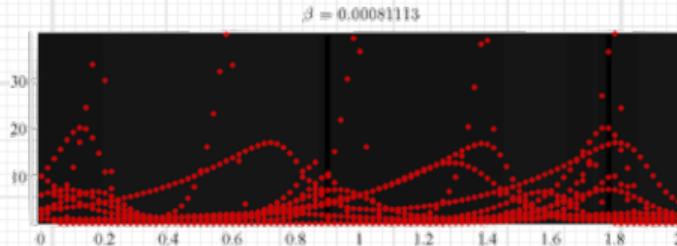
$$y' = x_1 y - x_2 y^2 \quad \text{unknown parameter } \mathbf{x} = [x_1, x_2]^\top$$



## ④ Lotka-Volterra

$$\begin{aligned}y'_1 &= x_1 y_1 - x_2 y_1 y_2 \\y'_2 &= -x_3 y_2 + x_4 y_1 y_2\end{aligned}$$

$$\mathbf{x} = [x_1, x_2, x_3, x_4]^\top$$



$p_j^{\max}$

$p_j = 0$

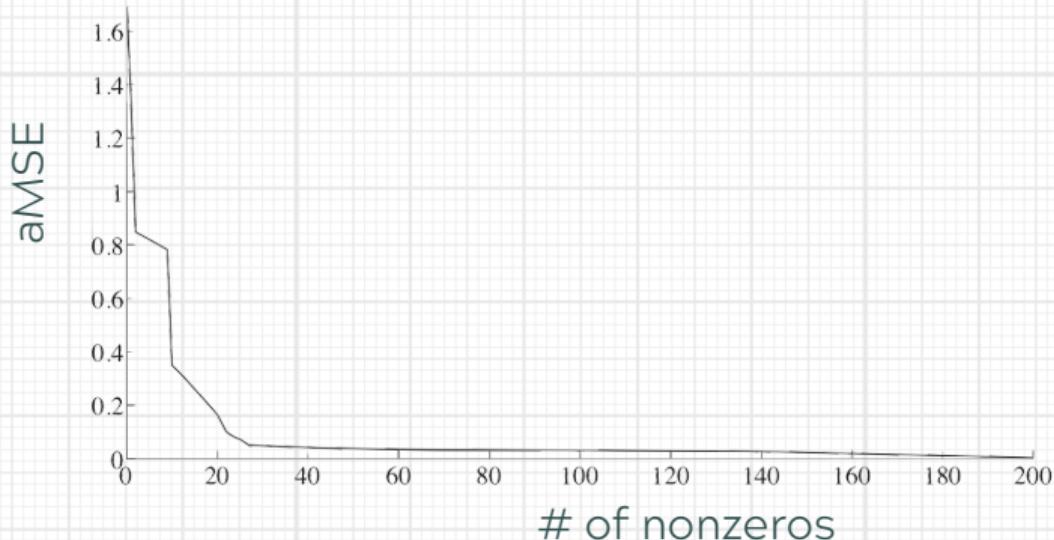
samples = 10, discretization 202, no noise, and  $\mathbf{L} = 0.1 \mathbf{I}_n$

## ④ Lotka-Volterra

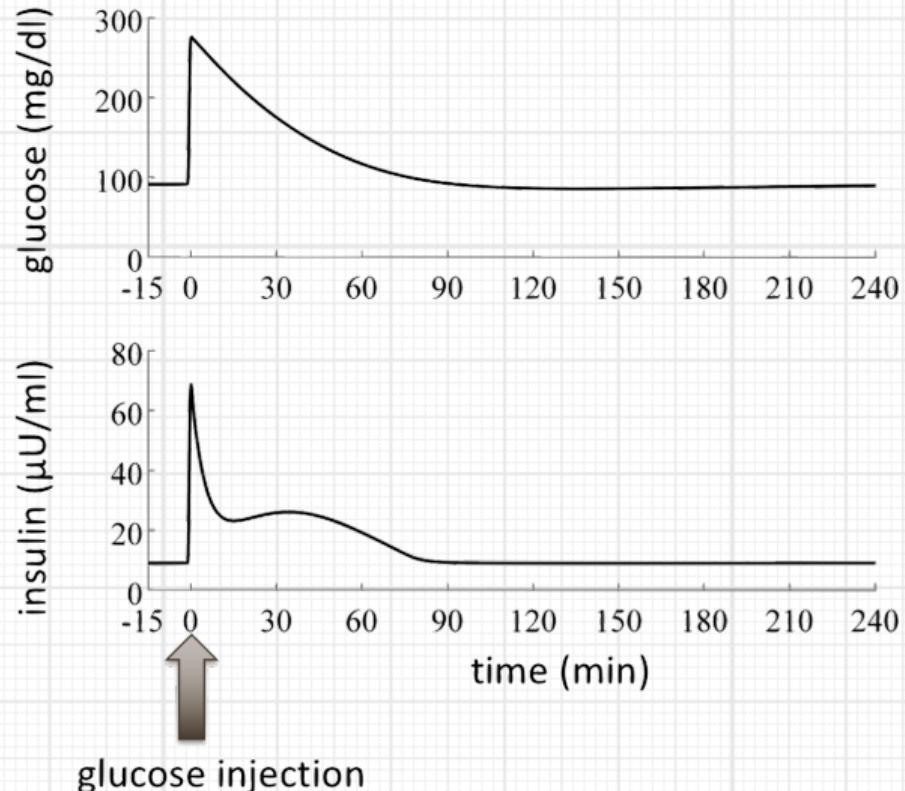
$$y'_1 = x_1 y_1 - x_2 y_1 y_2$$

$$y'_2 = -x_3 y_2 + x_4 y_1 y_2$$

$$\mathbf{x} = [x_1, x_2, x_3, x_4]^\top$$



## ④ intravenous glucose tolerance test (IVGTT)



## ④ glucose minimal model

### Minimal Model

$$\dot{G}(t) = -x_1 + X(t)G(t) + x_1 G_b$$

$$\dot{I}(t) = -\gamma \max(G(t) - h, 0)t - n(I(t) - I_b)$$

$$\dot{X}(t) = -x_2 X(t) + x_3(I(t) - I_b)$$

$G, I, X$  blood glucose, plasma insulin, effective insulin

$G_b, I_b$  basal level of glucose and insulin

$\gamma$  pancreatic insulin release rate

$h$  pancreatic threshold

$n$  degradation rate of insulin

$x_1$  glucose effectiveness

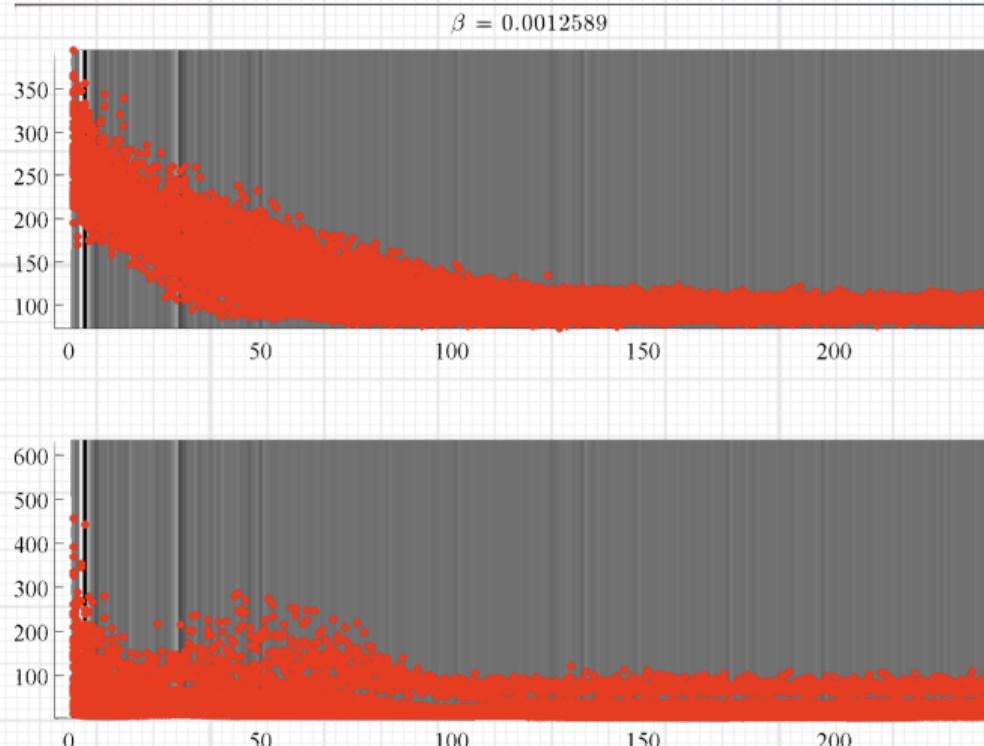
$x_2$  degradation rate of effective insulin

$x_3$  stimulation sensitivity of insulin

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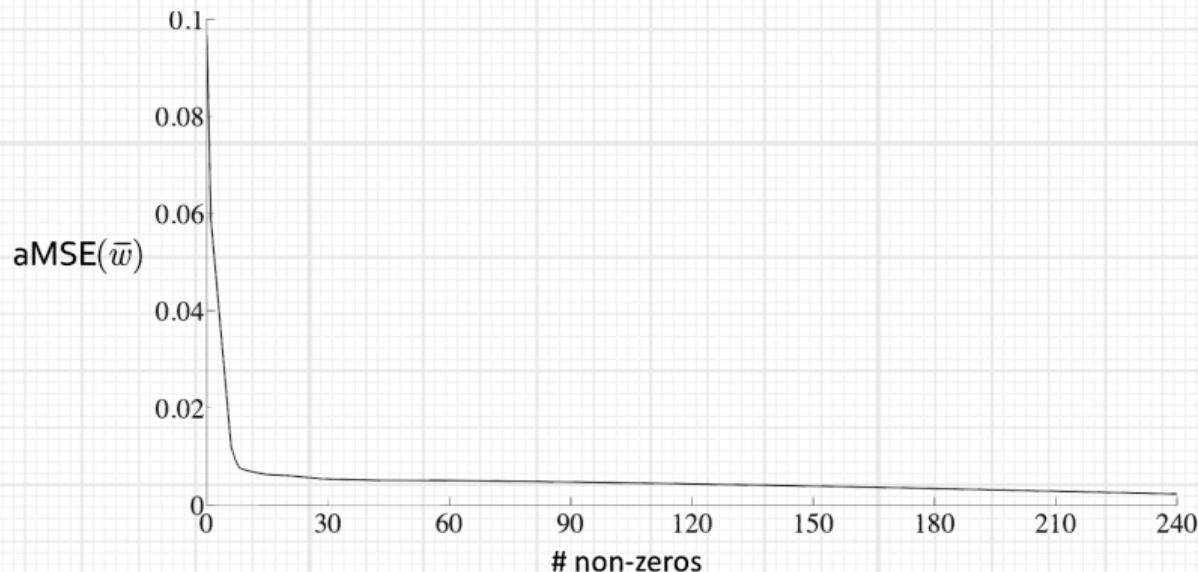
R N Bergman, L S Phillips, and C Cobelli. Physiologic evaluation of factors controlling glucose tolerance in man: measurement of insulin sensitivity and beta-cell glucose sensitivity from the response to intravenous glucose. In: *The Journal of clinical investigation* 68.6 (1981), pp. 1456–1467.

## ④ minimal model

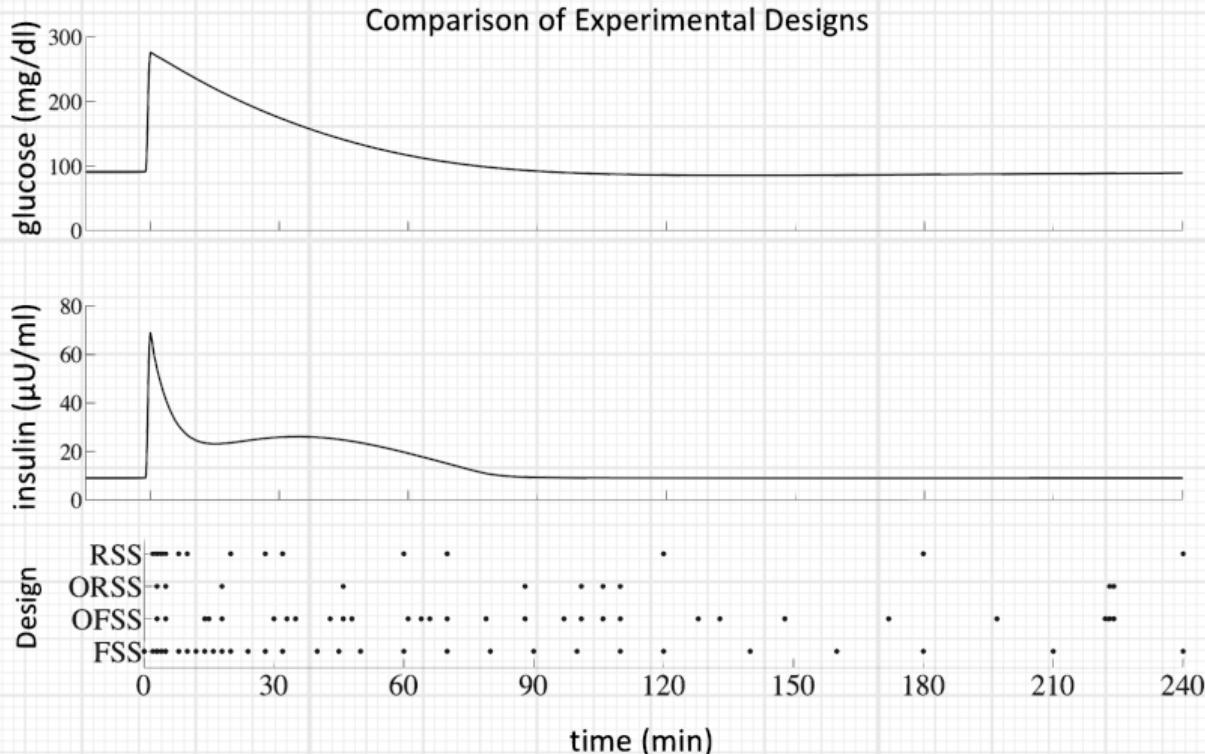


54 samples, discretization 241

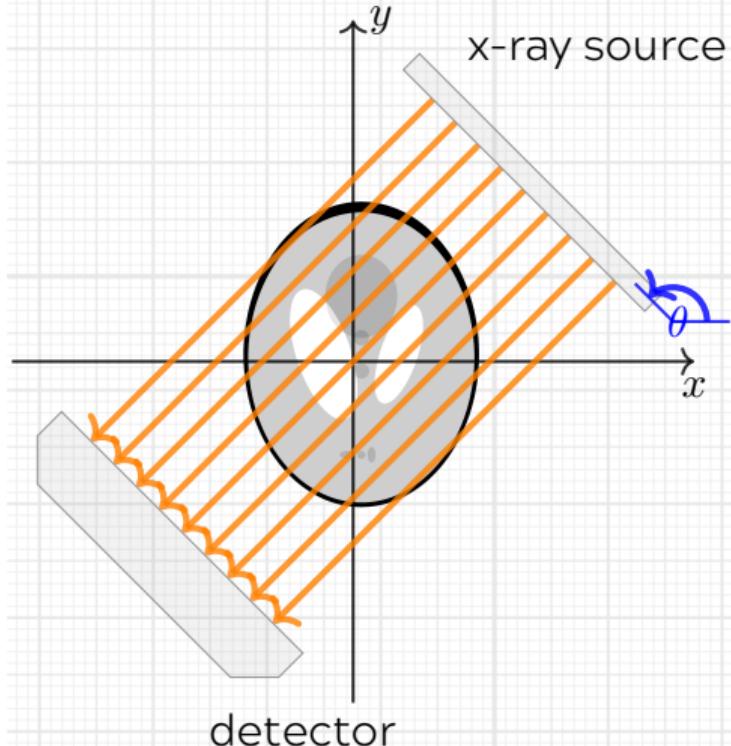
## ④ IVGTT: sparsity vs. error



## ④ IVGTT: proposed design



## ④ tomography



$$\mathbf{d}(\theta) = \mathbf{T}\mathbf{R}(\theta)\mathbf{x}_{\text{true}} + \boldsymbol{\varepsilon}(\theta)$$

- $\mathbf{x}_{\text{true}} \in \mathbb{R}^n$  true object
- $\theta$  angle
- $\mathbf{R}(\theta) \in \mathbb{R}^{n \times n}$  rotation of object
- $\mathbf{T} \in \mathbb{R}^{n_r \times n}$  transmission process

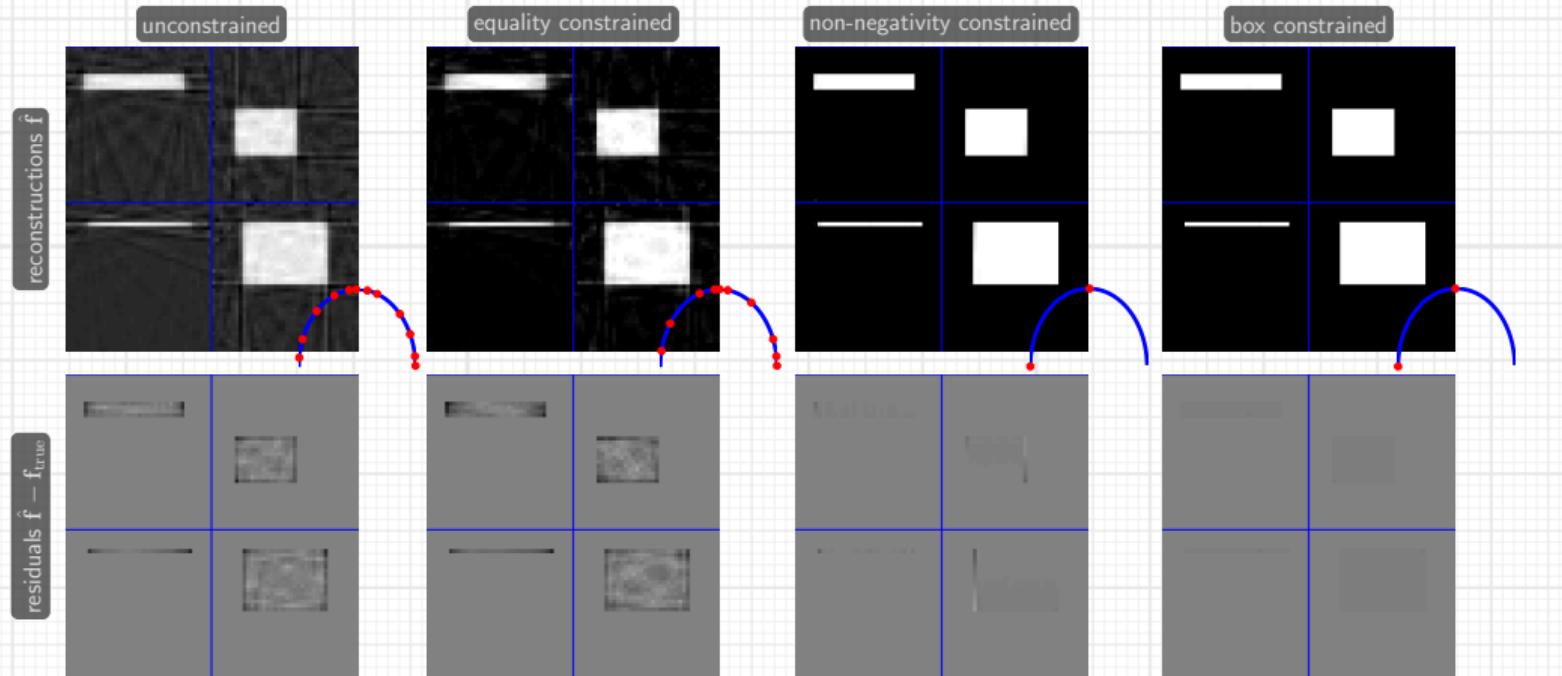
Where and how often should be measured for good recovery?

Design constraints

- |               |                     |
|---------------|---------------------|
| • time        | • cost              |
| • health risk | • limited resources |

L Ruthotto, J Chung, and M Chung. Optimal Experimental Design for Inverse Problems with State Constraints. In: *SIAM Journal on Scientific Computing* 40.4 (2018), B1080–B1100.

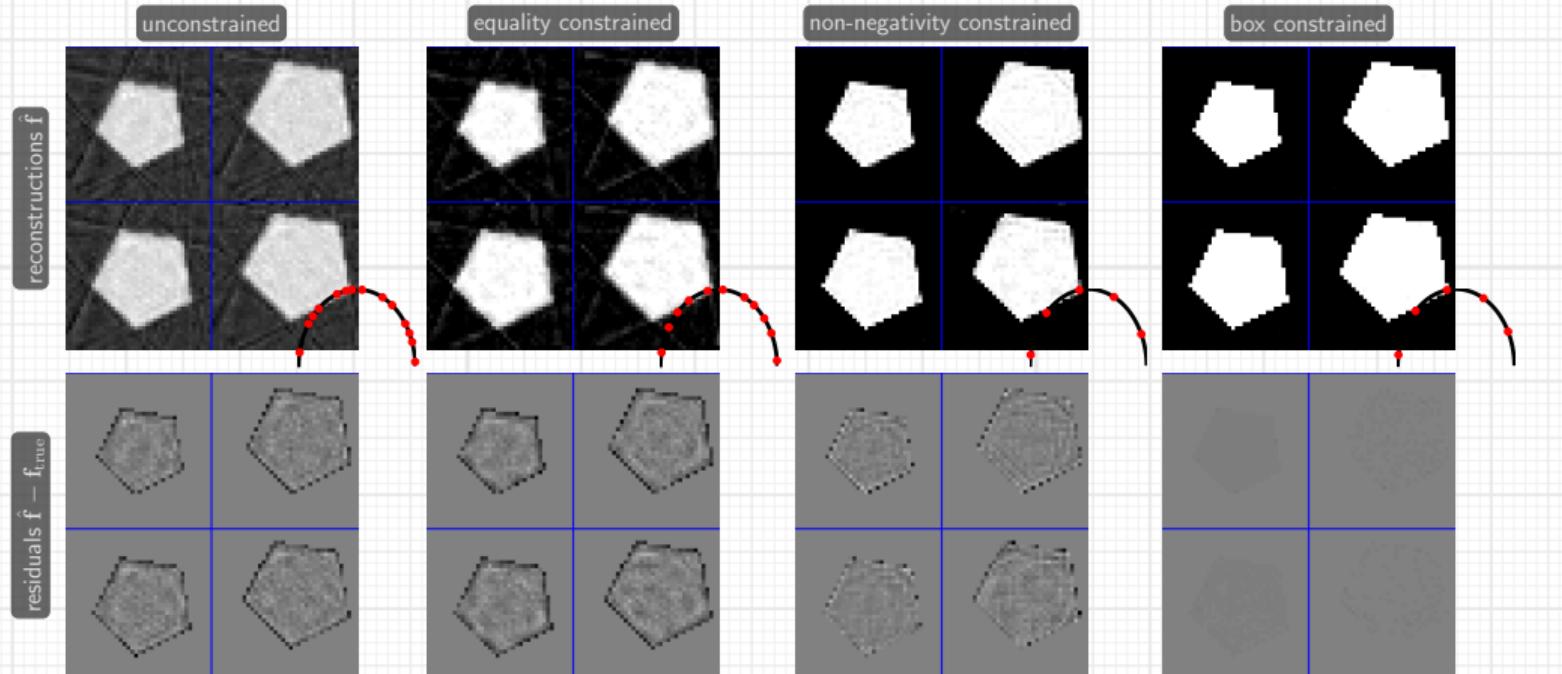
## ④ tomography results: rectangles



intuitive optimal angles [0, 90]

[0, 90] (non-negative)    [0, 90] (box)

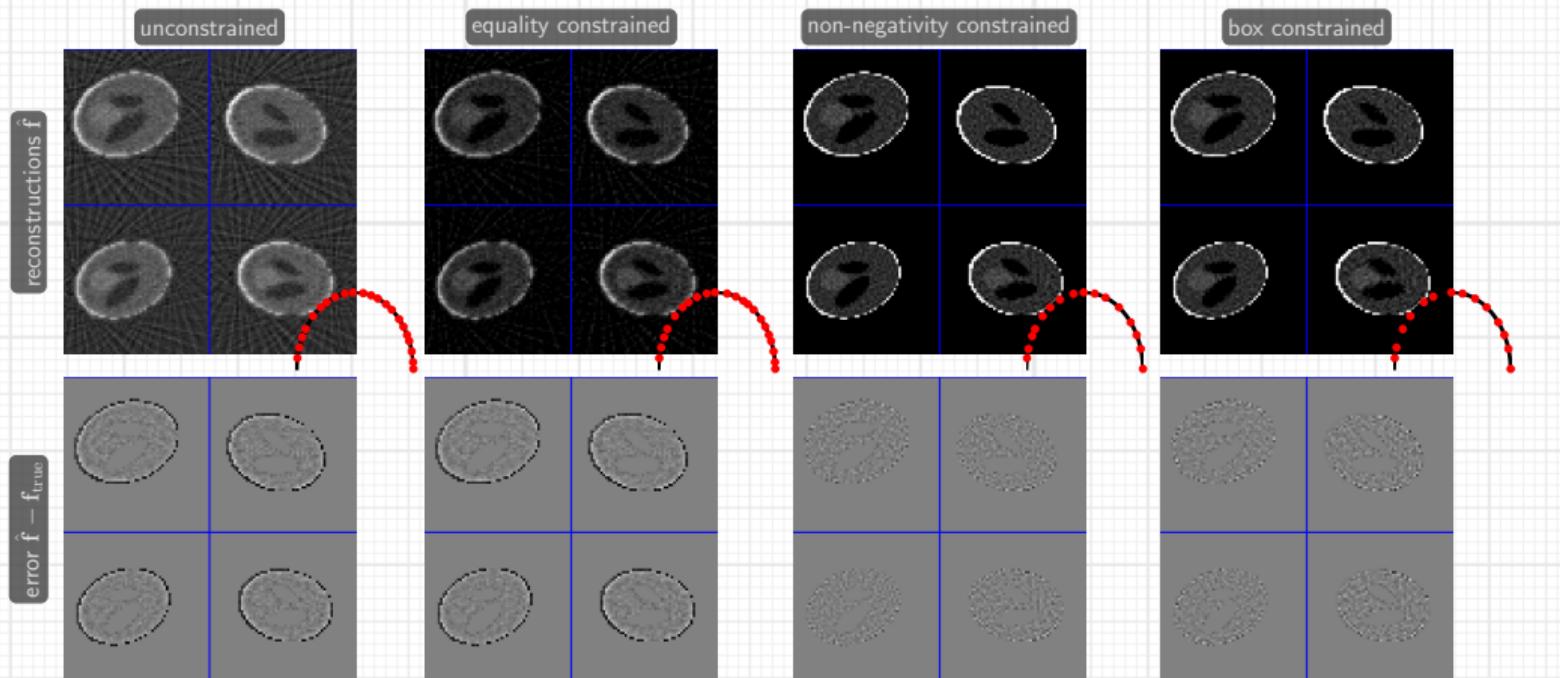
## ④ tomography results: pentagons



intuitive optimal angles [27, 63, 99, 135, 171]

[25, 64, 99, 136, 171] (non-negative)      [27, 62, 99, 134, 171] (box)

## ④ tomography results: Shepp-Logan phantom



## ④ conclusion & outlook

### Take-home message

- flexible and robust differential equation solvers
- efficient parameter estimation method using PDA
- robust parameter estimation via surrogate data
- new computational framework for optimal experimental design

### Outlook

- computational methods for finite element methods for ODE/DDE/DAE
- Gaussian processes for ODE PE, inverse problems, model reduction, missing data problems
- apply OED to various system setups

Thank you for your attention!

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